

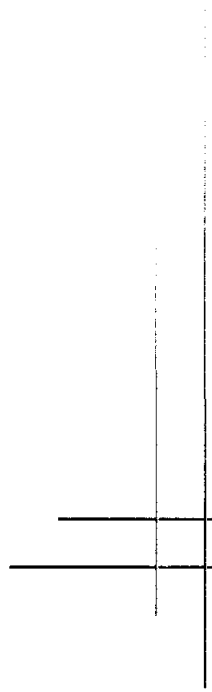

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SOUTHWEST
ECONOMETRICS, INC.

UNDERSTANDING TRENDS IN TEXAS
PER CAPITA WATER CONSUMPTION

Prepared for
Texas Water Development Board

by

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I. INTRODUCTION

Recent Trends in Per Capita Water Consumption

Water planning entities in Texas and other areas of the Southwest have long held a common expectation concerning the water use rates of urban communities, namely that per capita water consumption would continue its long established upward trend. Such expectations were reinforced year after year as the data became available on current water use from municipalities which record and report their water use to planning agencies. This experience of the data consistently reinforcing the expectation continued unabated until the mid to late 1970s when use rates stopped rising. During the 1980s per capita water use began to decline and seems to have established a long term reversal of the upward trend.

The data show that municipal per capita water consumption in Texas "increased from about 100 gallons per capita per day (gpcd) in the post World War II era to levels slightly above 182 gpcd by the mid-1970s. Subsequent to then, average per capita use in the State had leveled out and...in 1978 averaged about 178 gpcd. By 1987,...consumption had fallen to about 170 gpcd, exhibiting a general declining trend over the ten-year period..." (see **Figure I-1**).¹

These downward trends can be seen graphically in the data for several cities in Texas including Austin, San Antonio, Corpus Christi, Beaumont, Arlington and Pasadena (**Figure I-2**). These downward trends have major implications for water planners, especially since the planning horizons are very long in the discipline, reaching out some 40 years into the future in order to allow time for facility construction that often requires years of planning, permit processing, land acquisitions and construction.

Due to the importance of this long term trend to water policy and planning agencies, it is very essential to know the factors which are driving the downward trend in consumption rates.

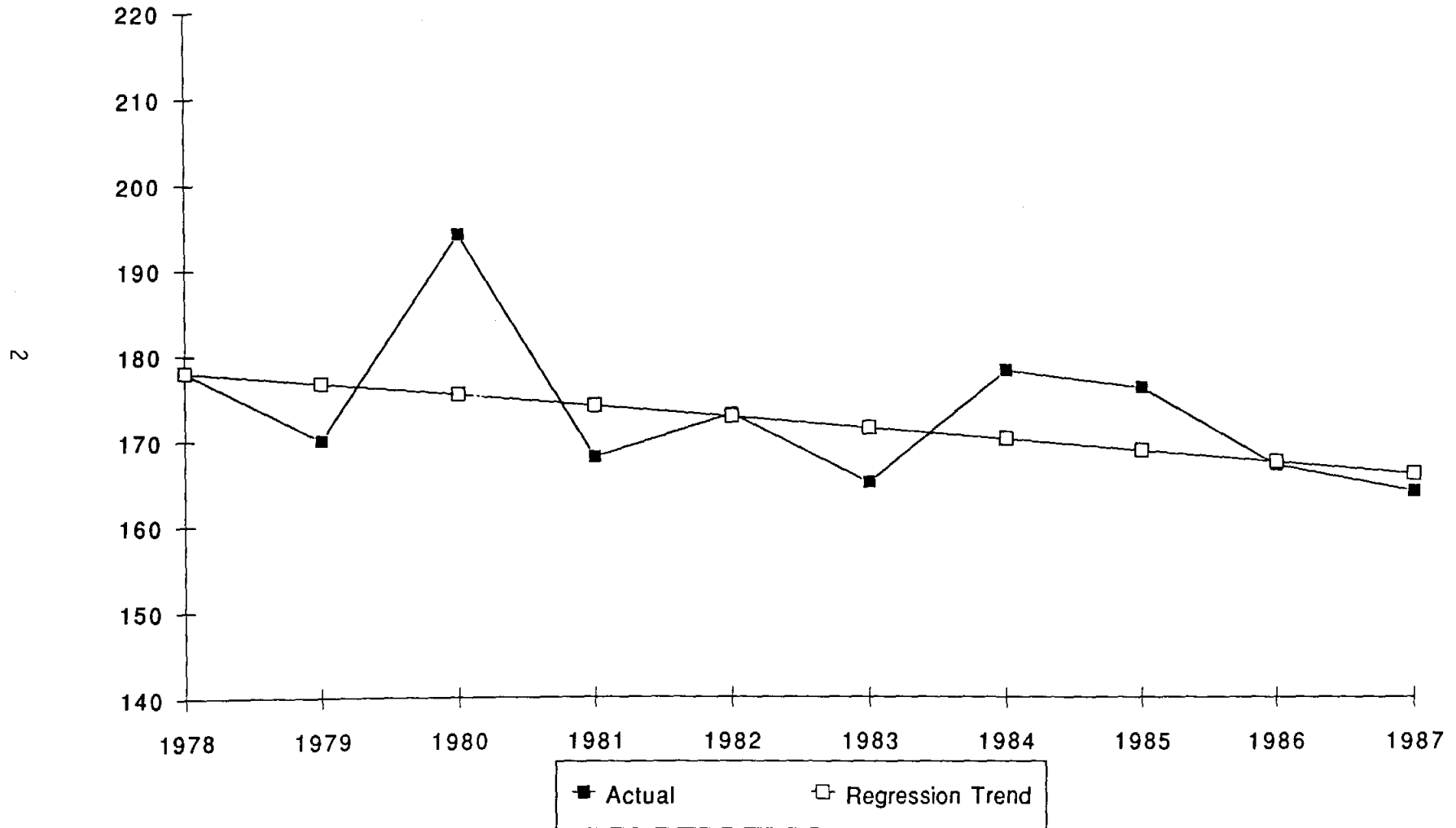
¹Water for Texas: Today and Tomorrow - 1990, published and distributed by the Texas Water Development Board, Austin, Texas, December, 1990, p. 2-9.

FIGURE I-1.

Texas Per Capita Municipal Water Use Trends

STATE OF TEXAS

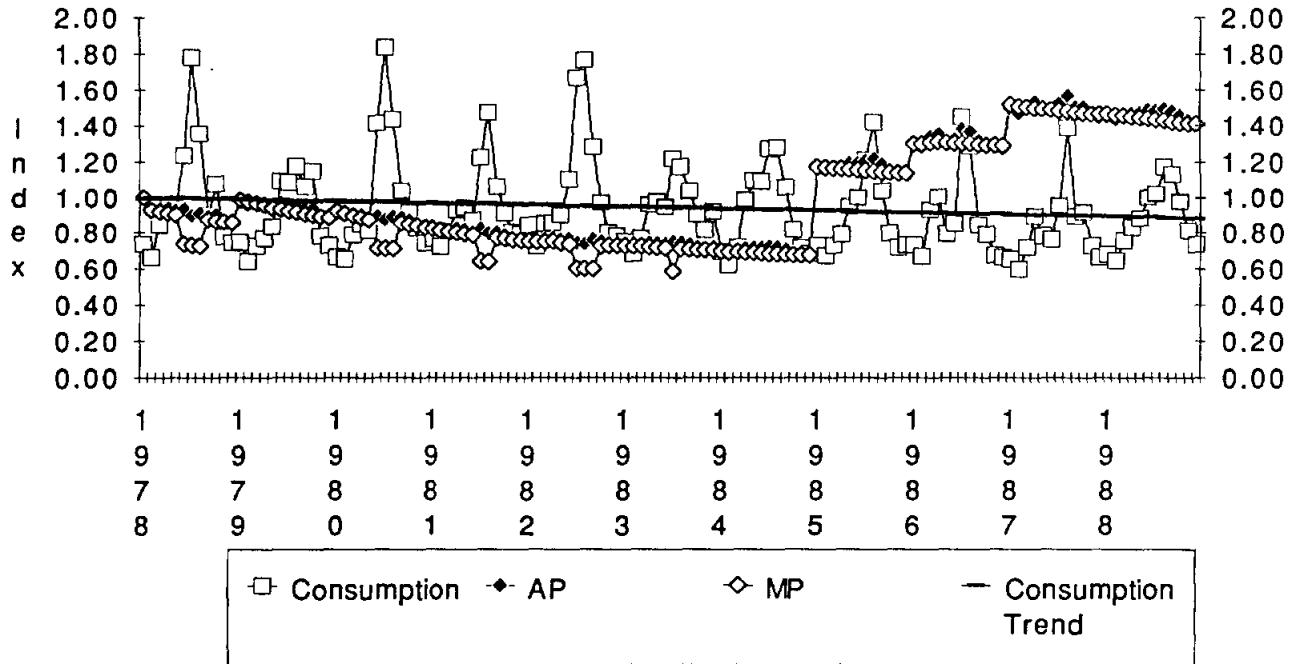
Avg. Gallons per Capita Daily



Source: Texas Water Development Board, Water for Texas: Today and Tomorrow - 1990, Austin, Texas, December 1990, p. 2-9.

Figure I-2. Trends in Per Capita Water Consumption in Major Texas Cities

Austin



Arlington

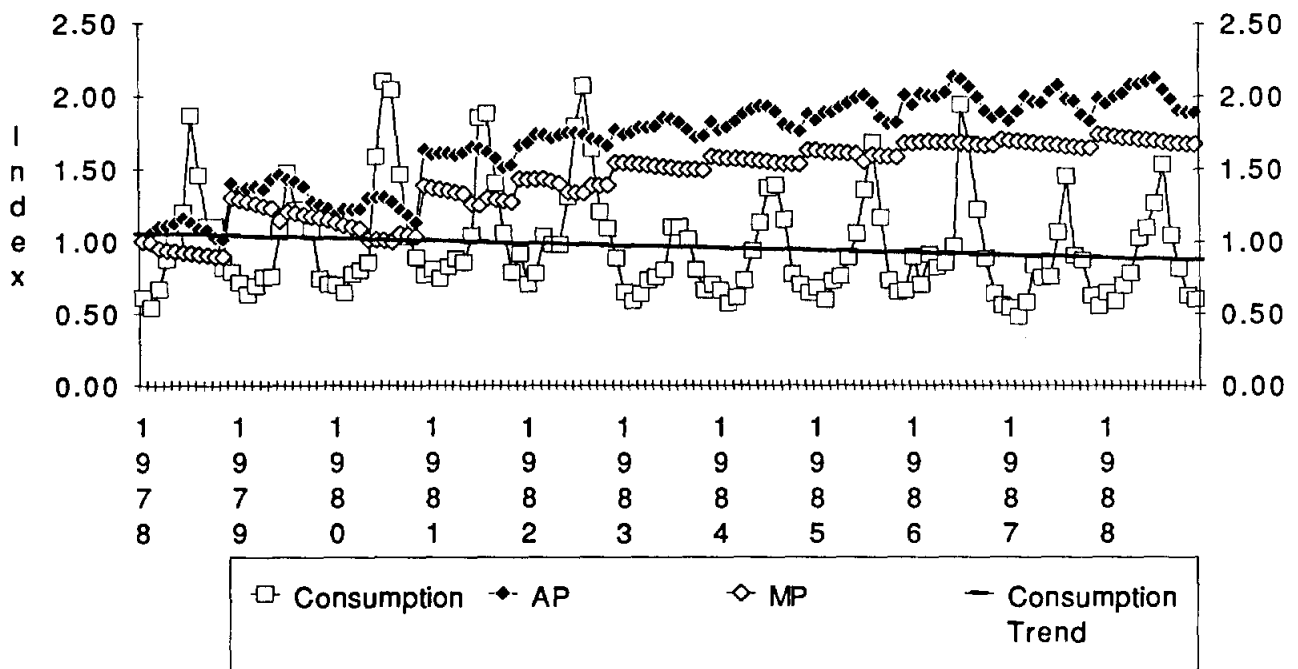
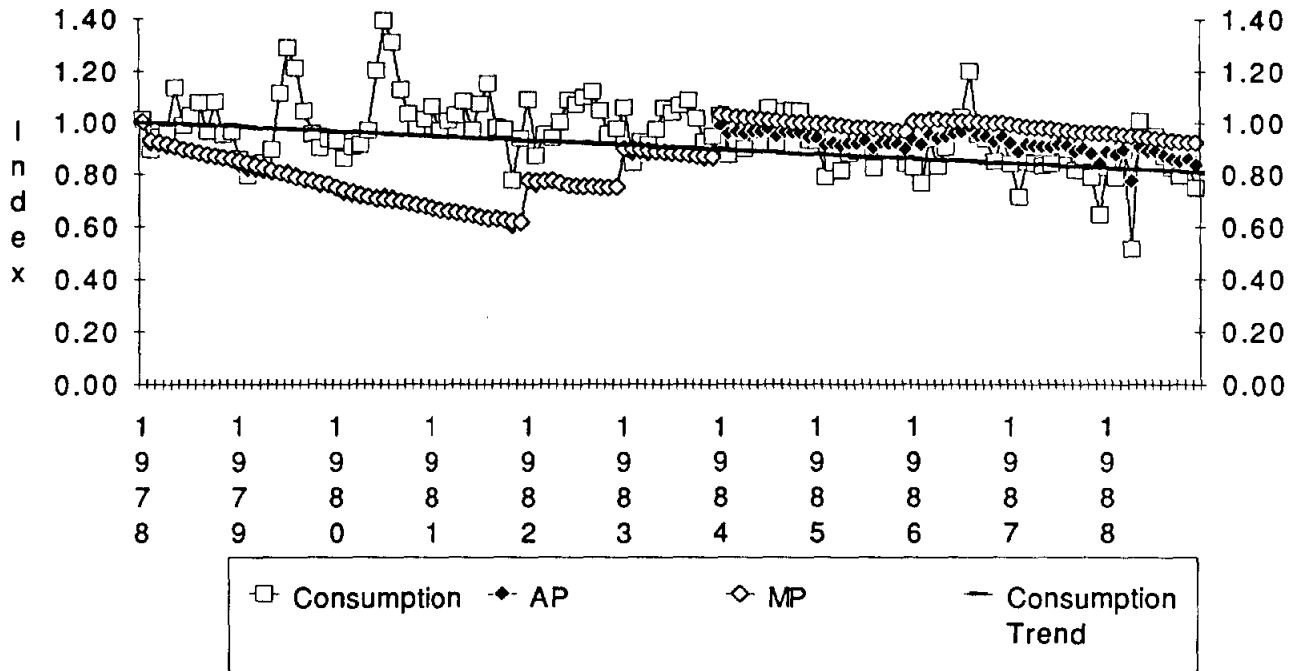


Figure I-2 (continued)

Beaumont



Corpus Christi

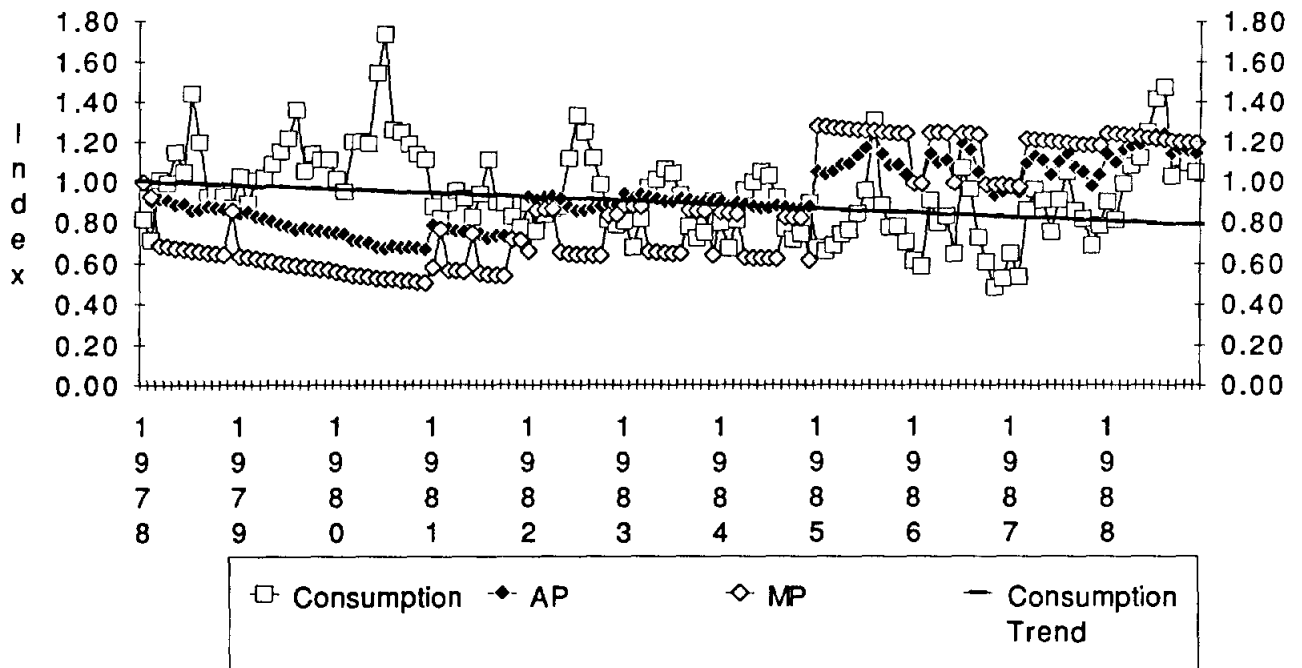
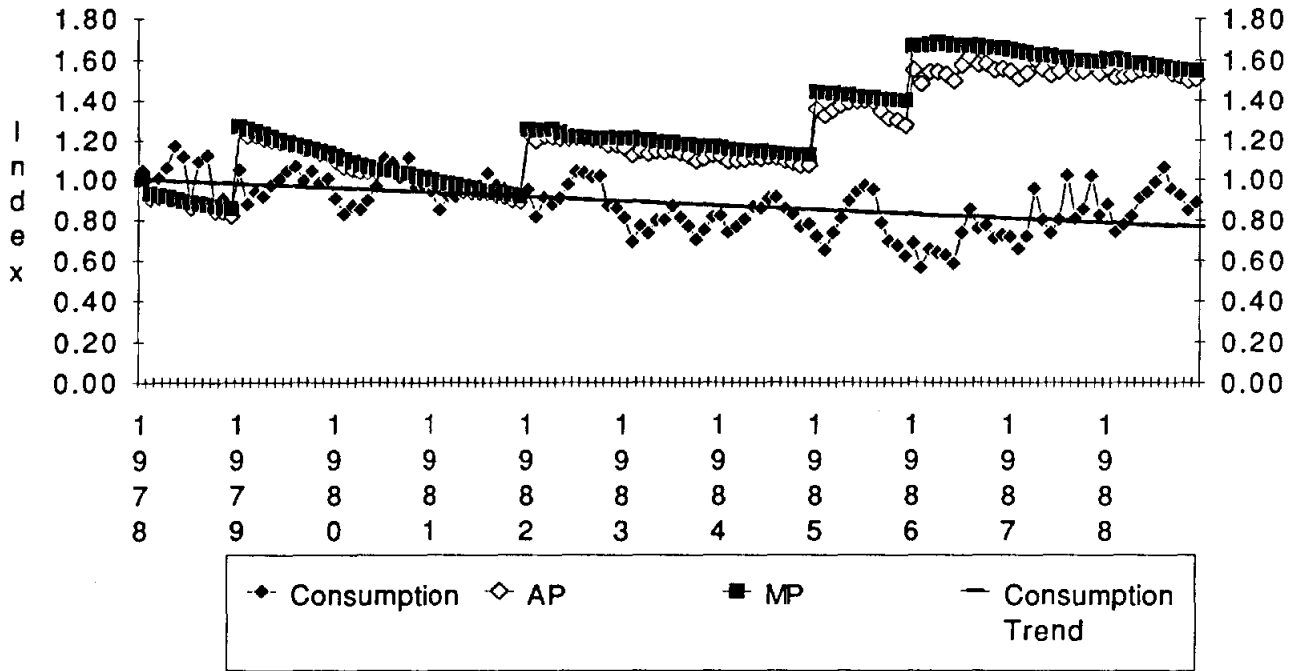
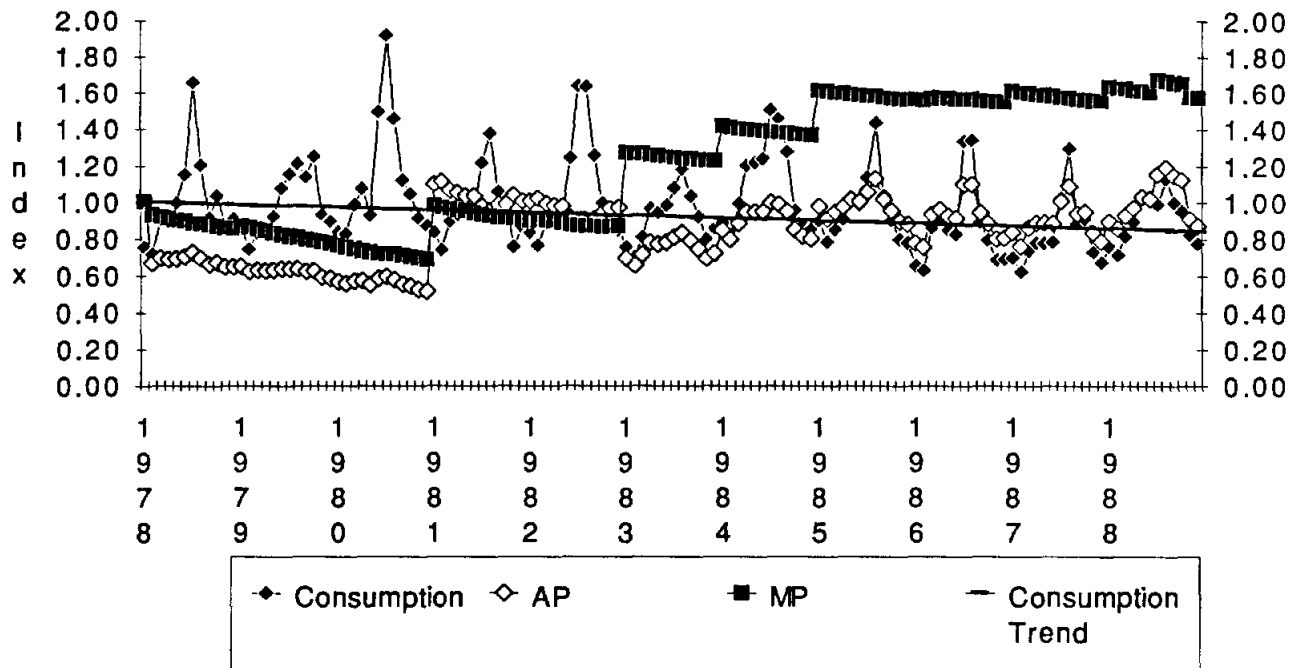


Figure I-2 (continued)

Pasadena



San Antonio



Price Structure Changes, Public Policy and Conservation

Several major events and public policy changes have, no doubt, had an important influence on water consumption in Texas including (1) the cyclical growth pattern of the Texas economy, (2) public policy in water and wastewater resources and (3) the development of a conservation ethic. During the 1970s and 1980s, Texas (and the nation) suffered three oil price shocks that brought long term shifts in the price structure of the economy which, in turn, have led to changes in energy costs of all types, changes in the use of materials and changes in the size of housing. These oil price shocks first increased the incomes of Texans relative to that of the nation in the 1970s and drastically reduced such in the 1980s following the oil price collapse in 1986.

The rapid growth of Texas communities in the late 1970s and early 1980s caused most communities to overbuild water and wastewater facilities since they expected growth to continue. Such overbuilding led to cost and rate increases above the long term trends.

Another important change that occurred during the period of the late 1970s and early 1980s was a major shift in Federal wastewater policy, namely the drastic reduction in funding for wastewater treatment plants, plants which were mandated to be built in order to meet Federal clean water standards. This reduction in Federal funding that had for years, stayed at 90% of the cost of wastewater treatment plants shifted the cost burden to state and local agencies. The end result was a significant increase in rates for wastewater services in many Texas communities. Federal Clean Water and Safe Drinking Water Acts also imposed higher costs on utilities by increasing the standards for the provision of wastewater and water services.

Another major public policy change occurred at the Texas State government level. A new statute amending the Texas Water Code was adopted by the Legislature in 1985 which, among other things, established a new set of financing mechanisms to provide low cost financing of local water projects by the extension of the state's credit capacity and cost of money to local government agencies. As a requirement of obtaining such State assistance, local entities were required to develop and submit a conservation plan that needs to include certain characteristics established by rule of the Water Development Board and/or by State statute.

Another factor, which is difficult to quantify, but which may indeed explain some of the trends at work, is the development of a broad based "conservation" ethic concerning the use of natural resources. The series of oil crises,

water shortages and related general environmental awareness has no doubt been responsible for the development of this conservation ethic.

Purpose of This Study

The purpose of this study is to determine the several factors that underlie and explain the recent downward trends in per capita municipal water consumption in Texas. Further, the purpose of the study is to quantify the relationship between these various factors and per capita water use.

II. REVIEW OF LITERATURE ON WATER CONSUMPTION

Econometric Models of Consumption

Econometric methods have been used extensively over the last thirty or so years to estimate the relationship among various economic, climatic and sociological factors and water consumption. The literature of such studies is dominated by the work of economists and engineers.

The economics profession naturally thinks of the type of problem posed by this study in terms of micro economic theory, namely the supply and demand for a consumer item. Therefore, a review of the literature in economics publications will usually turn up a list of studies that concern attempts to estimate "demand functions" that relate water consumption to price, income, structure of the decision making unit (households), weather and climatic factors, and perhaps the structure of price.

A literature review of the topic of explanations of water consumption will also turn up numerous studies completed by engineers. Because engineers are less concerned with theoretical underpinnings than are economists, one often finds in these studies a process modeling approach that is more rigorous as a perceptive tool than as a descriptive tool. That is, the only ability to test the explanatory power of the model against historical experience in such models is by comparative statistics and visual inspection. Such processing models have their greater strength in organizing the informed judgment of the authors into a system that allows the inclusion of new influences not in the historical data. For example, one can model the expected influences of conservation programs on future water consumption by use of a process model even though conservation programs did not exist during the historical period from which economists derive data for econometric models. More discussion of this topic is included in the next section of the report.

The review of literature here focuses primarily on econometric models of water consumption since it is our main purpose to identify and quantify the factors that explain the recent trends in per capita water consumption. Because of the potential importance of conservation programs on the long term future of water consumption, however, some attention is also given to processing models and methods for considering influences in the future that do not exist in the historical experience.

A comprehensive review of the econometric water demand models was completed by the U.S. Army Corps of Engineers in

1984.² This publication not only reviews the work to date but focuses on the question of price elasticities of demand in water consumption. Several recent studies have been completed that focus directly on Texas.³

Boland et al. reviewed more than 50 substantial studies of the response of municipal and industrial water use to price. The review included mostly work done on the topic since the 1960s. Not only did Boland et al. review the content of the studies but made judgements about the statistical rigor of the studies and drew conclusions about the range of price elasticities that characterize demand by summer and winter use by region of the U.S. by user class. The studies typically included explanatory variables of price, number of households, persons per household, household income, property value, irrigable area and climate.

Boland et al. found that there had been a number of conventions used in the specification of the consumption variable and the price variable, as well as others. Important in explaining the overall variation in consumption in residential use is the number of households, and while often a statistically significant variable, price makes a relatively small contribution to the overall explanation of the variance in consumption.

Results of the Boland et al. study indicate that price elasticities of demand for water are likely to be in the following ranges:

²Boland, John, Bondedykt Dziegielewski, Duane Baumann and Eva Opitz, Planning and Management Consultants, Ltd., **Influence of Price and Rate Structures on Municipal and Industrial Water Use**, for U.S. Army Corps of Engineers, Institute for Water Resources, Carbondale, IL, June 1984.

³The most comprehensive studies done in Texas in recent years include two state-wide studies by Ronald Griffin at Texas A&M and a focus study of the Texas Mexican border water demand by Milton Holloway. The Griffin studies are: Griffin, Ronald C., and Chan Chang, **Community Water Demand in Texas**, Texas Water Resources Institute, Texas A&M University, April 1989, and "Community Water Demand: New Specifications," **Western Agricultural Economics Association**, Honolulu, Hawaii, July 10-12, 1988. The Holloway study is: Holloway, Milton L. and Doug Tharp, **A Methodology for Determining Ability to Pay: For Use in the Implementation of the Economically Distressed Areas Water Assistance Program**, for the Texas Water Development Board, Austin, Texas, March 1990.

| | Elasticity | |
|--------------------------|-----------------|------------------|
| | <u>Long Run</u> | <u>Short Run</u> |
| Residential (winter) | 0.0 to -0.10 | NA |
| Residential (summer) | | |
| Eastern U.S. | -0.50 to -0.60 | NA |
| Residential (sprinkling) | | |
| Eastern U.S. | -1.30 to -1.60 | NA |
| Western U.S. | -0.70 to -0.90 | NA |
| Residential Average | -0.20 to -0.40 | 0.0 to -0.30 |
| Commercial | -0.20 to -1.40 | NA |

Griffin and Chang, in the 1989 study, estimated demand functions for municipal water use in Texas communities. The data base included 221 communities with data for the period 1981-1985.⁴ Griffin found that average price is empirically preferred to marginal price, and that the monthly price elasticities are on the order of -.14 in the winter and -0.28 to -0.37 in the summer, measured at the means of monthly consumption. Griffin tested prices in both real and nominal terms, but did not express a strong preference as to which to use. He also included a sewer price in the definition of water price and found that it should be included. That is, he performed a test of sorts and concluded that the data suggest that consumers don't know water and wastewater prices separately, or individually for that matter, but instead are aware of only the monthly bill. The monthly bill usually includes both water and wastewater.

The Holloway and Tharp study of 1990 had the purpose of estimating the ability of communities in the Texas/Mexico border area to pay for water and wastewater services based on what persons of similar economic circumstances were in fact paying. The modeling involved the estimation of regional demand functions from cross-sectional data derived mostly from the Bureau of the Census 1980 public use sample data. Price elasticity estimates are in the neighborhood of -0.80 for the communities included in this study.

Measuring Conservation Effects

There is naturally a strong interest of policy makers and planners in knowing the extent to which conservation programs of one type or another have any impact on water consumption. It seems clear, for example, that an intensive public awareness program during a drought period, or perhaps the initiation of an odd-even day watering scheme, has a

⁴Several major MSA cities were eliminated from the data set, however, including San Angelo, Plano, Pasadena, Mesquite, McAllen, Lubbock, Houston, Grand Prairie, Dallas, Ft. Worth, Baytown and Abilene.

significant impact on water use. It is much less clear whether such programs of awareness or regulatory restriction have any lasting impact beyond the current crisis. On the other hand, incentive programs for the installation of low-flow shower heads and toilet dams are sure to have a measurable, long term impact on non-sprinkling water use.

A statistically valid method of evaluation of the contribution of conservation programs would be to complete a controlled experiment where selected households would be identified for record keeping over a long period of time. The group would be offered the benefits of a particular conservation program, e.g. free low-flow shower heads and toilet blocks. Those who took advantage of the program would be put in one class and those who did not, in another. Records on consumption, price, income, household size, weather and climate would be maintained over a long enough period of time to determine the behavior of the group who elected to take advantage of the program. One would want to know, for example, whether the shower heads were replaced with regular models and whether the units stayed in place after the house was sold, etc. That is, a well designed test of the difference made by the conservation program would require some control over the data in order to apply normal statistical methods to the question. There are no such studies in the literature that we are aware of.

Another method of analysis would be applied at the community level rather than the individual consuming level. That is, it is possible to statistically compare the water use of communities over time that have, among all the variables that tend to explain consumption, a set of communities that have conservation programs with those who do not. There is no evidence in the literature that such a study has been completed.

III. METHODOLOGY FOR EXPLAINING PER CAPITA WATER CONSUMPTION TRENDS

The method of analysis for explaining the Texas trends in per capita water use selected for this study is that of multiple regression. Time series and cross-sectional data were combined in a data base of per capita consumption for 72 communities for the time period of 1978-1988. The communities were grouped into various regional groupings and equations were estimated for each. The equations allow one to explain recent trends in consumption and to use these equations for forecasting future consumption. The data base and model specification are explained, respectively, in Chapters IV and V.

Definitions

There have been a number of conventions developed from studies of water consumption that have been considered in developing the current study. Some have been accepted and others have not. The issues surrounding the selection of variables, geographical coverage and time period of analysis are discussed in this section. Specific definitions of variables are in Chapter IV.

Consumption

Three types of consumption data are typically found in the literature on water demand. "Metered data" are most often used since they are unique to the consumer decision making unit (household or business) and are readily available from utility records. A second type is that of survey results where consumption is a derived calculation based on "reported expenditures" by the survey respondent. A third type is a "calculated disappearance" quantity that may be derived from gross withdrawal data reported by a utility serving a community.

The metered data has the obvious strength of being directly derived from the behavior of the decision making unit, the consumer, who makes the choices of budgeting and purchasing that is of the greatest interest in demand analyses. One weakness of this data is that billings information is collected by address, not household or business, so that the behavior of the decision maker over a period of time is not preserved in the data. Second, if one is interested in the explanation of, and forecast of, aggregate water consumption for a city, county or state, individual billing data are massive amounts of information to manage. Still another, and perhaps the most important weakness of billings data, is the absence of associated income, household size, housing characteristics or other

likely independent variables for use in econometric analyses.

Survey results that allow derivation of consumption via billings and rate structure information have the strengths of going directly to the decision making unit to gather the data on consumption, and at the same time gathering income, household size, housing characteristics, etc. that are needed for econometric analysis. One weakness of such data is that respondents do not usually know quantity consumed and may be able to provide only "ball park" expenditure information from which to derive consumption. A second weakness is that of cost. If one is interested in explaining and forecasting community and state level water consumption, survey data are expensive to obtain and often impractical to gather.

Gross withdrawal derived data have the strength of accounting for the total water consumption of a community. One of the weaknesses is that it is impossible to capture the direct association of the individual decision maker's water consumption with income, household size, housing characteristics, etc. that are needed to distinguish water consumption behavior in the context of budgeting and consuming decisions. Therefore, this type of data only allow analyses among communities where each community is, in essence, treated as a decision making unit. This allows the consumption of water to be associated with income, household size, climate and weather factors and price at the community level.

The consumption data selected for this study are a mixture of two of the three types discussed above. The Water Development Board has for years, maintained a community level data base of water consumption consisting of derived annual per capita consumption and a monthly distribution function that allows the derivation of per capita consumption by month by community. This data base is constructed by first calculating the total disappearance of water within a community based on the net of gross withdrawals and wholesale sales of the utility that serves the community. Since large industrial users are usually independent of the utility serving the community, the resulting data are residential plus small and medium commercial users divided by the number of people in the community. While this characterization of "per capita water consumption" has some obvious weakness, such as the variance due to the number and character of commercial consumers in the data, from an overall perspective, these data are the best available for this study. First, the data base represents the combined experience of several professionals over a long period of time with knowledge of each community in the set. Second, the data are based on utility reported information that captures accurately the total water

consumption by month for each community. For purposes of explaining and forecasting community level consumption for the State of Texas, this is the preferred data base.

Price

The problem of how to characterize the price of water for demand analyses is not an easy one. First of all, consumers usually are faced with a combined bill for water and wastewater so that price for the individual service of providing water and wastewater on a monthly basis is not recognized by the typical consumer. Second, there is the problem of whether one is interested in the average or the marginal price which in today's utility pricing are often quite different. Another problem is a conceptual one for demand analysis purposes. Theoretically one expects to use the marginal price for demand analysis because the micro economic theory of markets tells us that prices are determined at the margin or, said another way, that individual consumers are always faced by the marginal price when deciding whether or how much additional service or commodity to purchase.

There are many practical problems that cloud this issue. It is practically difficult to obtain either average or marginal prices for individual consumers since the only access to such information is through surveys or individual billings data. Community averages can be derived from total consumption and total utility revenues, but the marginal price that corresponds to that average must incorporate the rate structure of the utility.

The data chosen for the study are described in detail in Chapter IV, but as a general matter it was determined that we should test both the average and marginal price and that such could be derived by combining the average monthly per capita consumption data with average number of persons per connection and the rate structure for the utility, for each year of the period 1978-1988. That is, average and marginal prices at the average per capita consumption level by month by community can be derived for each community in the data set.

Income

A significant explanatory variable for explaining the level of water consumption among communities and over time is income. This variable captures a combined set of factors that relate to the housing and commercial building stock in a community that, in turn, has much to do with water consumption. The idea is that income of a community determines the size and character of housing for the residential sector and building space for the commercial sector. Implicit in the purchase of such building space is

the number of bathrooms and showers, as well as the size of lawn which requires irrigation. As income rises over time, water use and water using capacity also tend to rise, other things equal.

The only comprehensively available source for income data is that of the Bureau of Economic Analysis, U.S. Department of Commerce. This data is available at the county level for years up through 1988. This data was selected to represent each community within each county of the data set.

Weather

Any study seeking to explain variations in water consumption will need to take account of weather conditions that have an obvious influence on short term variations in consumption. Weather conditions may also influence consumption over the long term within the cycles of weather patterns that sometimes last for years.

Fortunately, Texas has a large number of weather stations located throughout the State, such that there is a data gathering system near almost every city within the Metropolitan Statistical Areas (MSA) of Texas. These data are public information and available through organized data systems such as the Texas Natural Resources Information System (TNRIS).

The variables of interest from weather station sources include temperature and precipitation. The expectation is that summer sprinkling water use, in particular, is heavily influenced by the extent of hot, dry days when transpiration rates are high. While forecasting by use of equations estimated from historical data will normally assume normal long term weather conditions, it is essential that adequate weather representations be included in the use of econometric models of historical consumption. Typical representation of weather includes maximum or average daily temperature and number of dry days. The specific form of these variables for the current study are discussed in Chapter IV.

Conservation

The term "conservation" has a variety of popular uses, but there is no commonly accepted definition of the term for analytical work. One confusing area concerns the use of the rate structure for "conservation" purposes. If asked to list and describe the conservation programs being implemented, utility employees often list a change in rate structure that has been revised from a declining or flat price per unit to an increasing block structure. One might well classify the use of an increasing block rate structure

as a conservation program if the increasing structure, in fact, bears no correspondence with marginal costs. The matter is complicated by conventions of pricing by regulated utilities.

Other common responses to questions about the definition of conservation programs include education, rationing during drought conditions, and subsidy programs to encourage investment in water savings technologies such as low-flow shower heads and toilet dams.

For our purposes in this study we have included all rate structures in the calculation of average and marginal prices for each city included in the analysis, regardless of whether the utility listed the rate structure as a conservation program. Education programs and subsidy programs to encourage installation of water conserving technology have been classified as conservation programs. Rationing during drought periods has been eliminated from the analysis altogether.

Conservation, as defined above, has been included in the analysis by testing whether there is a statistically discernable difference in per capita consumption due to the presence of a program. That is, a conservation variable has been defined and included in the regression analysis. The method of inclusion was to identify which communities have conservation programs and at what point in time they were begun.

Extent of Geographical Coverage

This study of per capita water use trends distinguishes municipal water use from other types of use, namely, industrial, electric utility and agricultural uses. The analysis deals only with municipal water use which includes the retail water sales of utilities to residential and small commercial classes of users. Large commercial/industrial users typically provide their own water and wastewater service or purchase from a utility on an individual contract basis.

Since the study is designed to deal with municipal water use only, a decision was made to limit the data gathering and analysis to the set of 72 cities that make up the 28 MSAs in Texas. These cities account for about 85% of the municipal water use in Texas.

Importance of Diversity Among Cities

One of the strengths of econometric analyses is that one can be definite about the population to which the analysis is applicable. That is, the statistical tests that

allow one to have confidence in the value of parameters estimated statistically apply only to the population from which the data were drawn. In this case, the results will apply to each and every city included in the analysis, but not to others. Therefore, the results will be strictly applicable to all the utilities within the MSAs of Texas that consume 85% of the municipal water in Texas; no more or no less.

The inclusion of each city in all of the MSAs in Texas insures that the diversity of climate, geology, culture and costs get considered. It insures that we will be able to derive meaningful information for the full range of diversity among Texas cities.

Changes Over Time

Most dynamic processes involve some time lapse before all of the influence of a prior change is fully played out. For example, if relative prices change, influencing the economics of choice between two consumer goods, the ability to take advantage of the favorable price change may involve some investment in new equipment, such that the change in consumption patterns is not really evident until later periods. For this and related reasons it is usually advisable to include several time periods in an analysis.

This study of per capita water use trends makes use of 11 years of data (1978-1988). One year completes a seasonal cycle of water use patterns that are influenced by weather and plant growing seasons. The inclusion of 11 years of time series data allows the analysis to span a major rise and fall of the Texas economy and to consider the lag effects that may accompany consumer response to price and income changes of the late 1970s and early 1980s.

Strengths and Weaknesses of Econometric Analyses

The strengths of econometric analyses fall into two categories. The first has to do with the degree of confidence one may have in the parameters that quantify relationships among variables. For example, the measure of consumer response to price changes can be estimated using such analyses, and these price effects may be separated from another influence such as income changes. Not only can one separate the effects of two such influences, but he can also derive statistical tests that allow a measure of confidence in the estimate.

Another strength of econometrics is that it allows one to analyze an enormous amount of data efficiently. The

current study, for example, involves 72 cities and 132 monthly observations each.

The only significant weakness of econometric analysis for the purposes of this project is that it is limited to factors and relationships that are present in the historical period. One cannot measure the future effects of a conservation program being put in place today using econometrics if there is no comparable set of programs included in the available historical data. Other modeling approaches will be required for such problems. For example, one could construct a prescriptive model based on cost minimizing behavioral assumptions for such a question.

Formulation of an Econometric Model of Texas Water Consumption

There are two sources for developing a hypothesized mathematical model of Texas municipal water consumption. One is from economic theory of consumer behavior that provides the information that consumers' economic choices tend to follow general rational responses, such as decreasing consumption when prices rise and increasing consumption when incomes rise (other things equal). From economic theory we bring the following information to the current problem:

- (1) quantity consumed is inversely related to price (inflation removed);
- (2) quantity consumed is directly related to incomes (inflation removed);
- (3) quantity consumed is directly related to family size or persons per household;
- (4) quantity consumed is directly related to temperature and plant moisture stress; and
- (5) there are complementary and substitute consumer products that may come into play when relative prices change.

Another important source of information for formulation of the current model of water consumption is the literature. A number of "hints" and "leads" come from past efforts to solve similar problems. From the literature, for example, we have expectations about the range of price elasticities that may come from the current work (see Chapter II). We also have accounts of variable definitions, mathematical formulations and statistical test results obtained by others.

Based on both sources of information discussed above, the following general model was formulated for the current analysis:

- Q_i = the per capita consumption in time period i ($i = 1 \dots 132$), for MSA j ($j = 1 \dots 28$)
- NP_{ij} = number of persons per connection in time period i and MSA j
- AP_{ij} = real average price per 1,000 gallons per month in time period i and MSA j
- MP_{ij} = real marginal price per 1,000 gallons per month in time period i and MSA j
- I_{ij} = real per capita income in time period i and MSA j
- T_{ij} = temperature by month in time period i and MSA j
- DD_{ij} = number of dry days per month in time period i and MSA j

Our prior expectation is that the signs on coefficients estimated for these variables will be positive (+) for NP, negative (-) for AP and MP, positive (+) for I, positive (+) for T and positive (+) for DD.

An alternative specification of a model would include a representation of a supply function and the equation specification would be a simultaneous equation set. Such a specification is needed, conceptually, to separate shifts over time in a demand function from movements along a demand function that may accompany supply function shifts. This problem in applied economics is known as the identification problem. The literature on the topic suggests that attempts to estimate a simultaneous equation set is unlikely to succeed. As discussed later, an attempt to estimate a simultaneous equation set did not prove successful here either.

IV. DATA BASE

Metropolitan Statistical Areas

Four types of data constitute the data base for evaluating trends in per capita municipal water use in Texas: water consumption, water price, income and climate. The Metropolitan Statistical Area (MSA) was the basic demographic unit by which these data were collected. The time period of record is 1978 through 1988.

Geographic Coverage

Cities (utilities) for which data were collected represent the cities within each MSA required to total at least 80 percent of the population of each of the 28 MSAs in Texas. This selection process yielded 72 cities with populations within a wide range (**Table IV-1**). Monthly data for 1978-1988 exist for each city in the data base.

Water Consumption

Monthly water consumption data were derived from the data base of the Water Development Board, which includes annual average per capita consumption per day, and the monthly distribution of the annual average daily consumption. The Water Development Board's population estimates and the number of residential connections reported by each utility were used to convert per capita daily consumption to household consumption in gallons per month. Household consumption is used instead of per capita consumption in order to be consistent with billing practices and to measure economic responses at the basic decision making unit.

Income

Income data are derived from the county per capita annual income estimates of the Bureau of Economics Analysis (BEA), regional Economic Measurement Division, as updated in May 1990. The BEA's income data were deflated by the consumer price index for the South (1982-84 = 100), published by the Bureau of Labor Statistics.

Average and Marginal Price

Average and marginal prices were derived for each city in the analysis through municipal rate schedules on file

TABLE IV-1. MSAs, CITIES AND COUNTIES

| MSA# | MSA | CITY# | UTILITY | COUNTY | MSA# | MSA | CITY# | UTILITY | COUNTY |
|------|-----------------------|-------|----------------------|-----------|------|-------------------|-------|---------------|------------|
| 1 | Abilene | 2 | Abilene | Taylor | 12 | Galveston | 219 | Friendswood | Galveston |
| 2 | Amarillo | 14 | Amarillo | Potter | 12 | Galveston | 227 | Galveston | Galveston |
| 3 | Austin | 30 | Austin | Travis | 12 | Galveston | 350 | League City | Galveston |
| 4 | Beaumont/PtArthur | 43 | Beaumont | Jefferson | 12 | Galveston | 602 | Texas City | Galveston |
| 4 | Beaumont/PtArthur | 476 | Port Arthur | Jefferson | 13 | Houston | 42 | Baytown | Harris |
| 5 | Brazoria | 13 | Alvin | Brazoria | 13 | Houston | 285 | Houston | Harris |
| 5 | Brazoria | 18 | Angleton | Brazoria | 13 | Houston | 456 | Pasadena | Harris |
| 5 | Brazoria | 72 | Brazoria | Brazoria | 13 | Houston | 130 | Conroe | Montgomery |
| 5 | Brazoria | 118 | Clute | Brazoria | 14 | Killeen/Temple | 322 | Killeen | Bell |
| 5 | Brazoria | 217 | Freeport | Brazoria | 14 | Killeen/Temple | 597 | Temple | Bell |
| 5 | Brazoria | 338 | Lake Jackson | Brazoria | 14 | Killeen/Temple | 134 | Copperas Cove | Coryell |
| 5 | Brazoria | 457 | Pearland | Brazoria | 15 | Laredo | 347 | Laredo | Webb |
| 6 | Brownsv/Harlingen | 80 | Brownsville | Cameron | 16 | Longview/Marshall | 321 | Kilgore | Gregg |
| 6 | Brownsv/Harlingen | 265 | Harlingen | Cameron | 16 | Longview/Marshall | 367 | Longview | Gregg |
| 7 | Bryan/College Station | 82 | Bryan | Brazos | 16 | Longview/Marshall | 388 | Marshall | Harrison |
| 7 | Bryan/College Station | 124 | College Station | Brazos | 17 | Lubbock | 370 | Lubbock | Lubbock |
| 8 | Corpus | 135 | Corpus | Nueces | 18 | McAl/Edin/Mission | 182 | Edinburg | Hidalgo |
| 9 | Dallas | 472 | Plano | Collin | 18 | McAl/Edin/Mission | 376 | McAllen | Hidalgo |
| 9 | Dallas | 98 | Carrollton | Dallas | 18 | McAl/Edin/Mission | 397 | Mercedes | Hidalgo |
| 9 | Dallas | 151 | Dallas | Dallas | 18 | McAl/Edin/Mission | 408 | Mission | Hidalgo |
| 9 | Dallas | 230 | Garland | Dallas | 18 | McAl/Edin/Mission | 463 | Pharr | Hidalgo |
| 9 | Dallas | 245 | Grand Prairie | Dallas | 18 | McAl/Edin/Mission | 638 | Weslaco | Hidalgo |
| 9 | Dallas | 298 | Irving | Dallas | 19 | Midland | 404 | Midland | Midland |
| 9 | Dallas | 401 | Mesquite | Dallas | 20 | Odessa | 438 | Odessa | Ector |
| 9 | Dallas | 498 | Richardson | Dallas | 21 | San Angelo | 529 | San Angelo | Tom Green |
| 9 | Dallas | 159 | Denton | Denton | 22 | San Antonio | 530 | San Antonio | Bexar |
| 10 | El Paso | 189 | El Paso | El Paso | 23 | Sherman/Denison | 158 | Denison | Grayson |
| 11 | Fort Worth | 115 | Cleburne | Johnson | 23 | Sherman/Denison | 556 | Sherman | Grayson |
| 11 | Fort Worth | 25 | Arlington | Tarrant | 24 | Texarkana | 429 | New Boston | Bowie |
| 11 | Fort Worth | 44 | Bedford | Tarrant | 24 | Texarkana | 601 | Texarkana | Bowie |
| 11 | Fort Worth | 193 | Euless | Tarrant | 24 | Texarkana | 628 | Wake Village | Bowie |
| 11 | Fort Worth | 213 | Fort Worth | Tarrant | 25 | Tyler | 613 | Tyler | Smith |
| 11 | Fort Worth | 249 | Grapevine | Tarrant | 26 | Victoria | 624 | Victoria | Victoria |
| 11 | Fort Worth | 261 | Haltom City | Tarrant | 27 | Waco | 47 | Bellmead | McLennan |
| 11 | Fort Worth | 293 | Hurst | Tarrant | 27 | Waco | 626 | Waco | McLennan |
| 11 | Fort Worth | 435 | North Richland Hills | Tarrant | 27 | Waco | 667 | Woodway | McLennan |
| | | | | | 28 | Wichita Falls | 654 | Wichita Falls | Wichita |

residential and commercial use. Weighted marginal and average prices were derived by the relative mix of residential and commercial water connections reported annually to the Water Development Board by each utility.

Climate

National Weather Service (NWS) data regarding precipitation and temperature for selected Texas weather stations were acquired from TNRIS. Data were selected from the NWS station nearest the city for which data exist for the period of record 1978-1988.

Temperature data are the average monthly temperatures at the NWS station nearest the city, and for which data exist for the period of record 1978-1988. Average monthly temperature is the mean of the average daily high temperature and average daily low temperatures as reported in two separate data bases at TNRIS.

Precipitation data are the total number of days in a month with less than 0.25 inches of precipitation at the NWS station nearest the city, and for which data exist for 1978-1988.

V. THE MODEL

Important Factors Affecting Consumption

The model for evaluating trends in per capita municipal water use in Texas was specified as a demand model, or a model in which the effect of price on consumption is measured. Important factors other than price which affect consumption are income, number of persons per household, average monthly temperature, the number of days per month without significant rainfall and the level of commercial development. Not all of these factors affect each city or MSA uniformly but are always important.

Nine Regional Models

Early analysis revealed that water consumption is better evaluated on a regional basis. Therefore, equations were estimated for nine regions of Texas. MSAs were grouped together based on a combination of criteria: location with respect to vegetational and geological designations, general precipitation patterns based on data from 1950 through 1981, commercial distinctiveness and city size.

Twenty-eight MSAs were grouped into the nine regions (**Table V-1**). The Metroplex (Dallas-Fort Worth) area constitutes two separate regions, each of which includes cities from both the Dallas MSA and the Fort Worth MSA. The distinction is due to suburban location of cities in both MSAs.

Cross-Sectional/Time-Series Combination

The period of record is 1978-1988 from which a data base was constructed containing monthly information for each city within each region. This time series affords the ability to analyze the response over time of water consumption to the explanatory variables of the model.

Grouping several MSAs, each of which contains one or more cities, allows cross-sectional analysis by which to examine the relationship between consumption and the explanatory variables for multiple locations within one period of time. The combination of time-series and cross-sectional data for analysis allows for explaining region-wide structural relationships and changes in those relationships over time.

TABLE V-1
MSA Groupings for Regional Models

| <u>REGION</u> | <u>MSA #</u> | <u>MSA</u> |
|---------------------------------|--------------|------------------------------|
| West | 10 | El Paso |
| | 20 | Odessa |
| | 19 | Midland |
| | 17 | Lubbock |
| | 2 | Amarillo |
| Rolling Plains | 28 | Wichita Falls |
| | 1 | Abilene |
| | 21 | San Angelo |
| | 15 | Laredo |
| Metroplex ¹ | 9 | Dallas |
| | 11 | Fort Worth |
| Metroplex Suburban ² | 9 | Dallas |
| | 11 | Fort Worth |
| Central | 23 | Sherman-Denison |
| | 27 | Waco |
| | 14 | Killeen-Temple |
| | 7 | Bryan-College Station |
| I-35 South | 3 | Austin |
| | 22 | San Antonio |
| | 26 | Victoria |
| Southeast | 5 | Brazoria |
| | 12 | Galveston |
| | 8 | Corpus Christi |
| East | 24 | Texarkana |
| | 25 | Tyler |
| | 16 | Longview-Marshall |
| | 4 | Beaumont-Port Arthur |
| | 13 | Houston |
| Valley | 6 | Brownsville- Harlingen |
| | 18 | McAllen-Edinburg- Mission |

¹Includes cities of Fort Worth, Arlington, Dallas, Plano, Carrollton, Irving and Richardson.

²Includes cities of Cleburne, Bedford, Euless, Grapevine, Haltom City, Hurst, North Richland Hills, Garland, Grand Prairie, Mesquite and Denton.

Functional Forms of Nine Regional Models

In general form, the model for water consumption is specified as:

$$\text{CONS} = f(\text{MP}, \text{FAMILYINC}, \text{TEMP}, \text{DAYS}, \text{COMPROXY}, \text{DSEAS}, \text{Dn})$$

where:

CONS = per household water consumption in gallons per month;

MP = weighted marginal price in dollars per thousand gallons;

FAMILYINC = per capita income multiplied by the number of persons per residential connection, in dollars;

TEMP = the average monthly temperature in degrees Fahrenheit;

DAYS = the number of days with precipitation of less than 0.25 inches;

COMPROXY = the fraction of total water connections attributable to commercial use;

DSEAS = dummy variable which distinguishes summertime consumption from consumption in the rest of the year.

Dn = locational dummy variables for MSA number n which distinguishes one MSA from another (number designations in **Table V-1**).

Three functional forms of each regional model were estimated econometrically using the Statistical Analysis System (SAS) regression procedure. The three forms are linear, log-linear and log-log, all results of which are in **Appendix A**.

Table V-2 contains the parameter estimates, t statistics, F test and number of observations (n) for the preferred functional forms for all nine regions. The log-linear form provided the best results for all regions except the East and the Rolling Plains. The Rolling Plains region was estimated in linear form. The East region was estimated in log-log form.

All parameter estimates in all nine equations are statistically significant with signs which are intuitively correct. The relationship between consumption and price is inverse, and the relationship between consumption and income is direct. The signs in the TEMP and DAYS parameters are all positive, indicating that higher monthly average temperatures and a larger number of days without significant rainfall tend to induce higher water consumption, all other variables remaining unchanged.

insignificant in early equations in which the sign on the parameter was negative.

Explanation of Historical Water Consumption

Data from seven cities from six separate regions were used to indicate the performance of the models in explaining historical consumption. **Figures V-1** through **V-7** show the actual vs. predicted values of dependent variables for water consumption in El Paso, Abilene, Dallas, San Antonio, Austin, Corpus Christi and Houston. Actual data are city-specific. The models are the respective regional models contained in **Table V-2**.

Variations in consumption were explained well for El Paso (**Figure V-1**), Abilene (**Figure V-2**), Dallas (**Figure V-3**), San Antonio (**Figure V-4**), Austin (**Figure V-5**) and Houston (**Figure V-7**) with predicted values approaching the actual values even in the summer peak consumption periods. Consumption patterns in El Paso, Abilene, Dallas, San Antonio, Austin, Corpus Christi and Houston appear to typify the patterns of their respective models. That is, the combined effects of variation in price, income, dry days, temperature and concentration of commercial water users do a good job of explaining the variation in monthly consumption in these cities using the applicable regional model. Only Corpus Christi (**Figure V-6**) reflects results atypical of its regional model.

The predicted values in Corpus Christi vary most from actual values in 1980, 1986 and 1988. Directional patterns show underestimation in 1980 and 1988, and overestimation in 1986. **Figure V-6** shows a tendency to over-predict in the last half of the period, which could be the result of the model's failure to capture the effect of conservation programs in Corpus implemented in the last half of the 1980s. The strong reversal to a high level of underestimation in 1988, however, seems to discount this possibility.

As mentioned earlier, attempts to estimate a conservation parameter did not prove successful. That is, the set of explanatory variables do not include conservation. Since a number of cities implemented conservation programs during the 1980s, we are interested in checking the patterns of predicted vs. actual consumption residuals to see if unexplained variation has a long-term trend that possibly could be explained by conservation programs.

An examination of residuals (**Appendix B**) indicates that the Southeast regional model explains variation in Lake Jackson and Galveston similar to the way it explains

Figure V-1

City of El Paso
Log of Consumption per Household

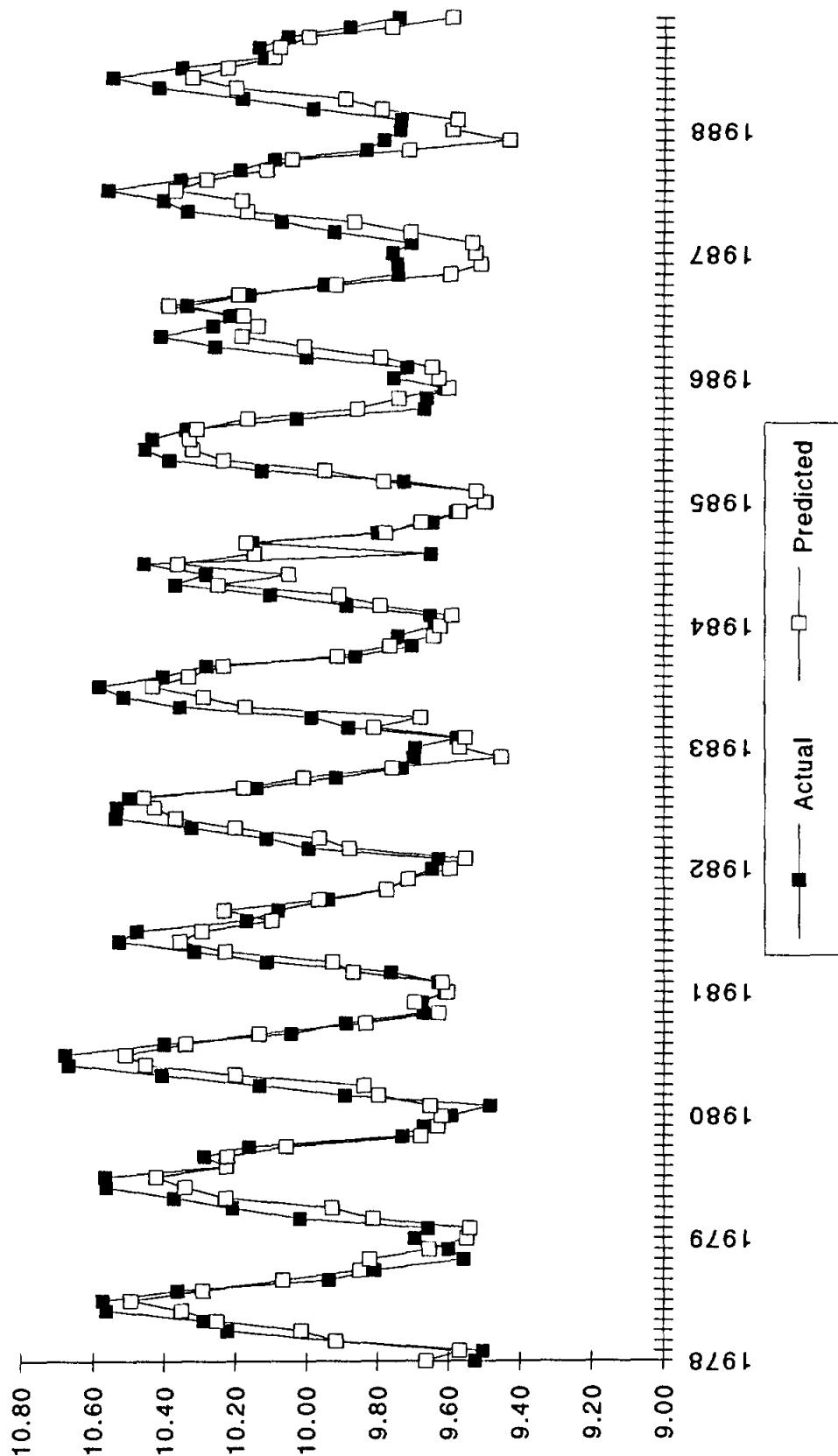


Figure V-2

City of Abilene Log of Consumption per Household

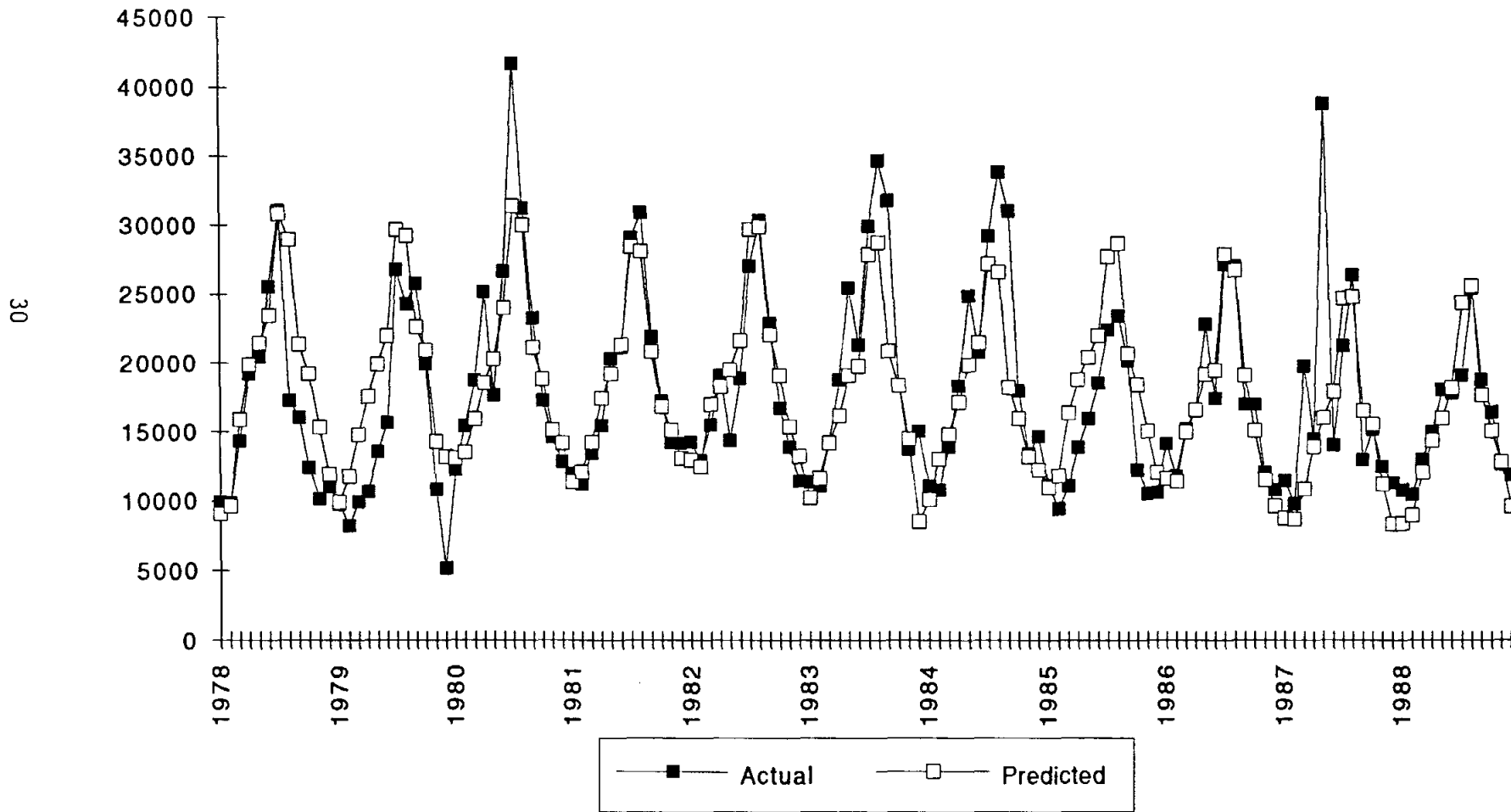


Figure V-4

City of San Antonio Log of Consumption per Household

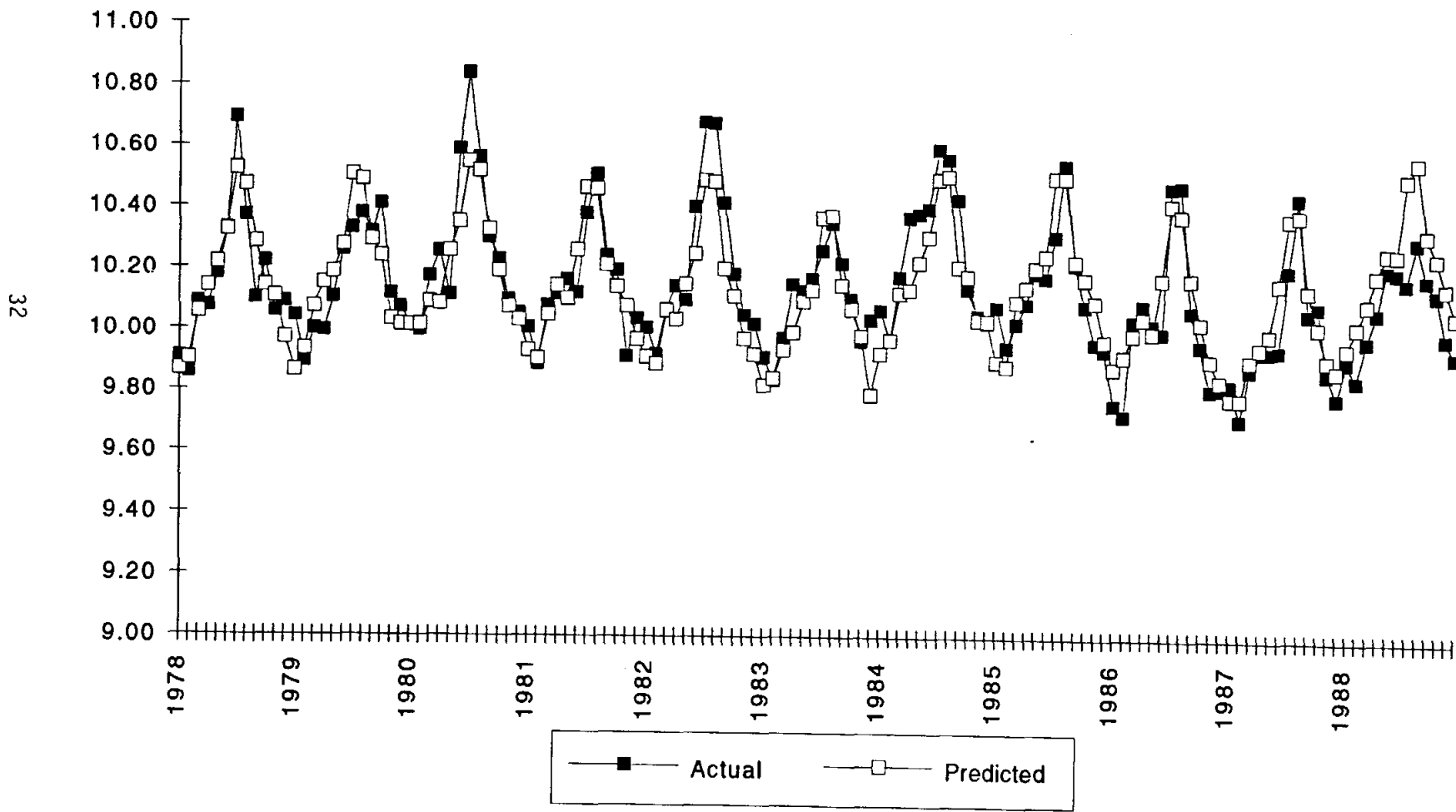


Figure V-5

City of Austin Log of Consumption per Household

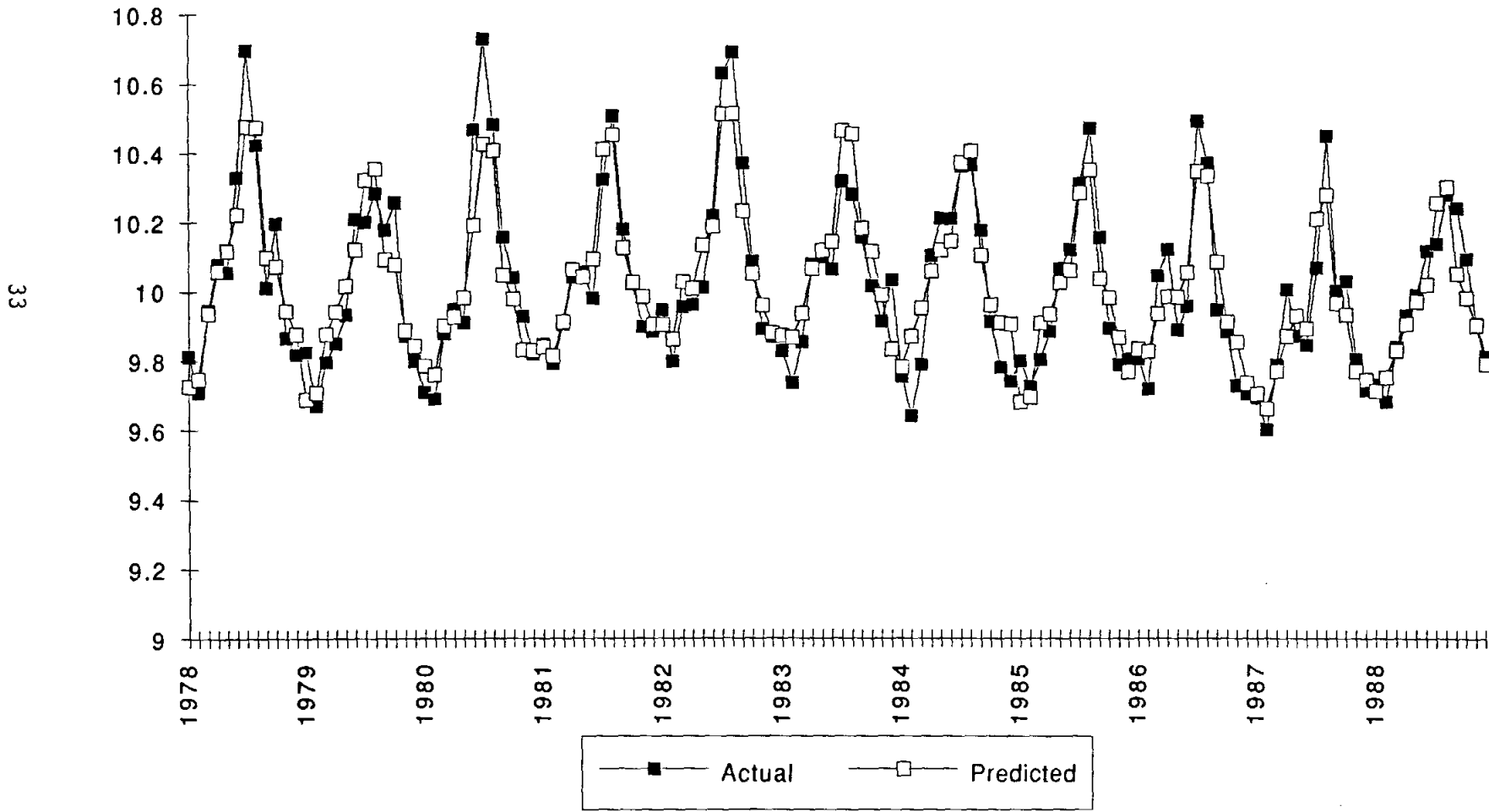


Figure V-6

City of Corpus Christi Log of Consumption per Household

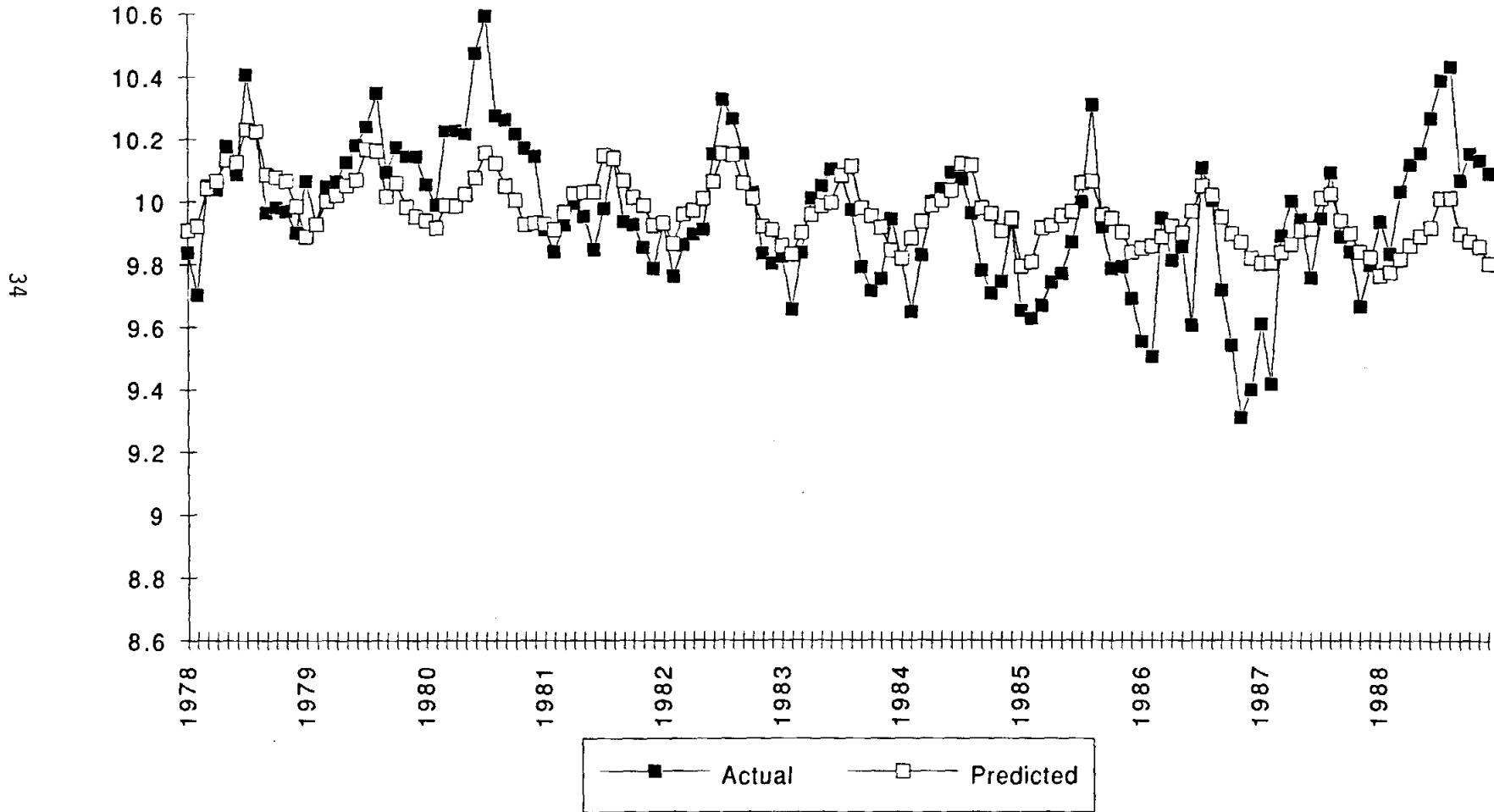
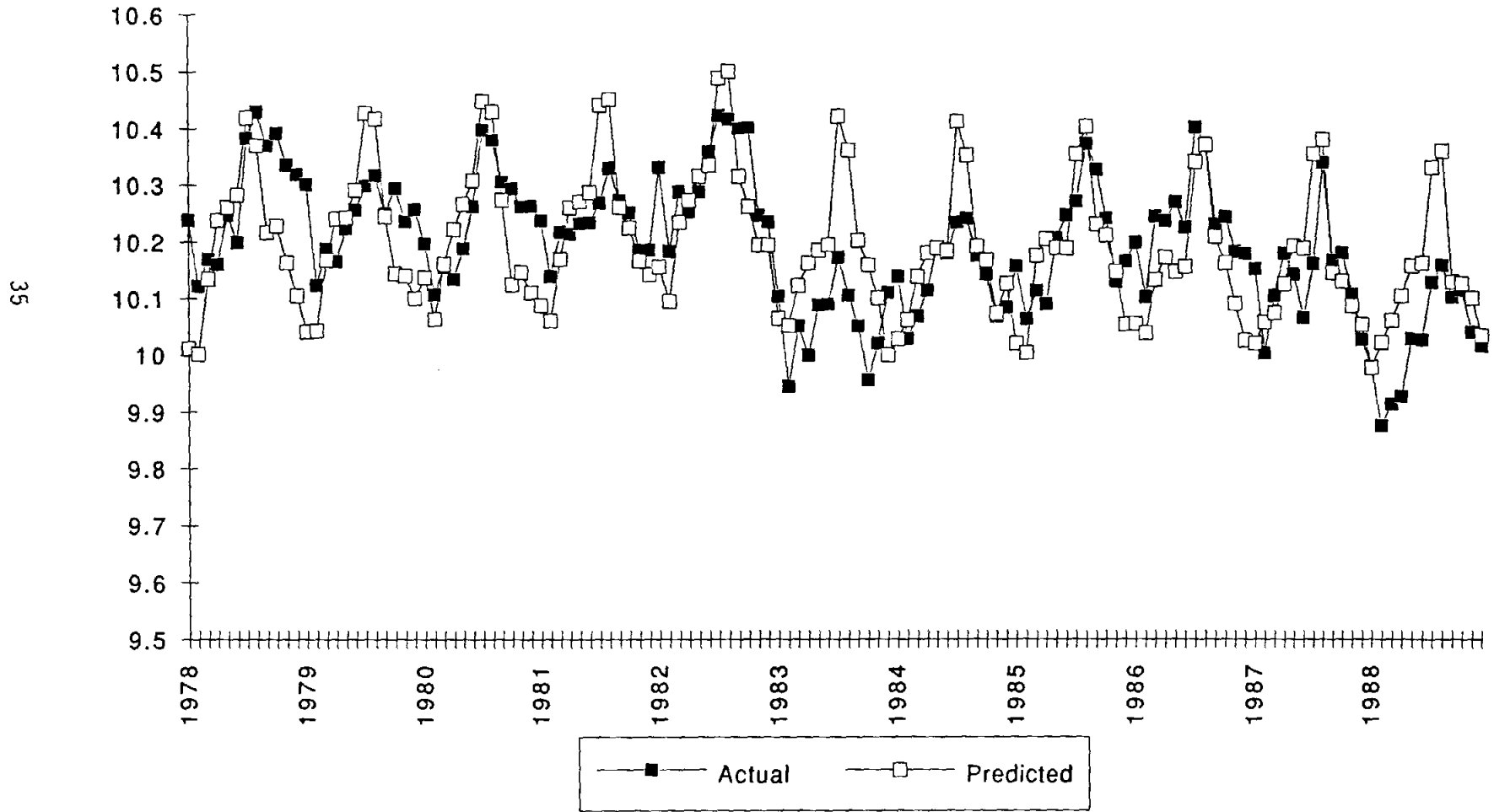


Figure V-7

City of Houston Log of Consumption per Household



actual consumption for all three cities, although Lake Jackson and Galveston had no conservation programs until as recently as 1988.

There is no constant pattern of overestimation at El Paso and Austin, two other Texas cities with notable conservation programs implemented in the mid 1980s. If overestimation in the late 1980s could be taken as an indication of the model's inability to capture the effects of conservation, residuals should show a positive upward trend. That is, the residuals would reflect this failure and show rising, positive values, but they do not (**Appendix B**).

In summary, the statistical tests of the models give us confidence that we have explained a large portion of the variation in monthly consumption by the set of variables that we expect to be important, namely price, income, commercial concentration, dry days, temperature and city/regional location. Further, we have confidence that while some equations show non-uniform patterns of actual minus predicted consumption, such patterns do not seem consistent with expected conservation program effects. While conservation programs are no doubt having some impact in certain cities, we are unable to quantify such with econometric methods, given the available data.

Elasticity Estimates

Price elasticity of demand measures the response of consumption to a one-percent change in the price of water. **Table V-3** shows summer and winter price elasticities for each region. The highest elasticity coefficients are for MSAs in the East region. The coefficients are lowest in the Valley region, indicating that consumption will drop only slightly as price rises. Water is a more precious commodity in the Valley than in the Rolling Plains. Price elasticity increases slightly for non-peak consumption periods except in Victoria (I-35), Dallas/Fort Worth (Metroplex), Sherman/Denison and Waco, where elasticity decreases in non-peak consumption periods, and in the East region which has constant elasticity coefficients due to the log-log form of the equation.

The response of consumption to a one-percent change in income is quite variable over the state (**Table V-4**). The response to rising income is lowest in the Valley region, at 0.031 in the McAllen-Edinburg-Mission MSA. The highest income elasticities are in MSAs in the Rolling Plains region, where Wichita Falls area residents tend to increase water consumption in winter by 2.3 percent for every one-percent increase in income.

TABLE V-3. PRICE ELASTICITY OF DEMAND FOR WATER BY REGION BY MSA IN TEXAS

| MSA\REGION | EAST | I-35 | METRO- PLEX | METRO- PLEX SUB | ROLLING PLAINS | WEST | SOUTH- CENTRAL EAST | VALLEY |
|------------|------------------|------------------|------------------|--------------------|-------------------|------------------|---------------------------|------------------|
| ABILENE | | | | | -0.173 -0.450 | | | |
| AMARILLO | | | | | | -0.047 -0.048 | | |
| AUSTIN | | -0.293 -0.316 | | | | | | |
| BEAUPORT | -0.090 -0.090 | | | | | | | |
| BRAZORIA | | | | | | | -0.087 -0.088 | |
| BRNSVHAR | | | | | | | | -0.024 -0.026 |
| BRYANCOL | | | | | | | | -0.127 -0.130 |
| CORPUS | | | | | | | -0.074 -0.078 | |
| DALLAS | | | -0.066 -0.065 | -0.177 -0.187 | | | | |
| EL PASO | | | | | | -0.042 -0.043 | | |
| FTWORTH | | | -0.066 -0.065 | -0.177 -0.187 | | | | |
| GALVESTON | | | | | | | -0.108 -0.110 | |
| HOUSTON | -0.090 -0.090 | | | | | | | |
| KILLTEMP | | | | | | | | -0.143 -0.147 |
| LAREDO | | | | | -0.095 -0.159 | | | |
| LONGMARS | -0.090 -0.090 | | | | | | | |
| LUBBOCK | | | | | | -0.068 -0.071 | | |
| MCAEDMIS | | | | | | | | -0.033 -0.034 |
| MIDLAND | | | | | | -0.075 -0.078 | | |
| ODESSA | | | | | | -0.072 -0.078 | | |
| SANGELO | | | | | -0.167 -0.288 | | | |
| SANTONIO | | -0.224 -0.228 | | | | | | |
| SHERMDEN | | | | | | | | -0.132 -0.130 |
| TEXARKNA | -0.090 -0.090 | | | | | | | |
| TYLER | -0.090 -0.090 | | | | | | | |
| VICTORIA | | -0.216 -0.207 | | | | | | |
| WACO | | | | | | | | -0.134 -0.133 |
| WICHITAF | | | | | -0.181 -0.543 | | | |

NOTE: The first elasticity in each set is the value for summer and the second is for winter.

TABLE V-4. INCOME ELASTICITY OF DEMAND FOR WATER BY REGION BY MSA IN TEXAS

| MSA\REGION | EAST | I-35 | METRO- PLEX | METRO- PLEX SUB | ROLLING PLAINS | WEST | SOUTH- CENTRAL EAST | VALLEY |
|------------|----------------|----------------|----------------|--------------------|-------------------|----------------|---------------------------|----------------|
| ABILENE | | | | | 0.818 2.056 | | | |
| AMARILLO | | | | | | 0.205 0.205 | | |
| AUSTIN | | 0.941 0.941 | | | | | | |
| BEAUPORT | 0.533 0.533 | | | | | | | |
| BRAZORIA | | | | | | | 0.738 0.738 | |
| BRNSVHAR | | | | | | | | 0.034 0.035 |
| BRYANCOL | | | | | | | | 1.267 1.267 |
| CORPUS | | | | | | | 0.712 0.712 | |
| DALLAS | | | 0.802 0.802 | 0.962 0.962 | | | | |
| EL PASO | | | | | | 0.194 0.194 | | |
| FTWORTH | | | 0.802 0.802 | 0.962 0.962 | | | | |
| GALVESTON | | | | | | | 0.733 0.733 | |
| HOLSTON | 0.533 0.533 | | | | | | | |
| KILLTEMP | | | | | | | | 0.978 0.978 |
| LAREDO | | | | | 0.492 0.812 | | | |
| LONGMARS | 0.533 0.533 | | | | | | | |
| LUBBOCK | | | | | | 0.244 0.244 | | |
| MCAEDMIS | | | | | | | | 0.031 0.032 |
| MIDLAND | | | | | | 0.302 0.302 | | |
| ODESSA | | | | | | 0.259 0.259 | | |
| SANGELO | | | | | 0.779 1.540 | | | |
| SANTONIO | | 0.948 0.948 | | | | | | |
| SHERMDEN | | | | | | | 1.006 1.006 | |
| TEXARKNA | 0.533 0.533 | | | | | | | |
| TYLER | 0.533 0.533 | | | | | | | |
| VICTORIA | | 0.840 0.840 | | | | | | |
| WACO | | | | | | | 1.015 1.015 | |
| WICHITAF | | | | | 0.836 2.301 | | | |

NOTE: The first elasticity in each set is the value for summer and the second is for winter.

VI. USE OF ECONOMETRIC MODELS FOR FORECASTING

This section of the report provides an example of how to use the econometric models for forecasting per capita monthly water consumption. The forecasts presented here are for exemplary purposes only. A forecast for planning use should be done paying particular attention to reasonable projections of region and city specific independent variables. The common set of assumed projections of independent variables used here are, however, within a reasonable range for the examples chosen.

Projections of Independent Variables

The independent variables which determine the following trends of per capita water consumption in Texas are price and income. All nine regional models include FAMLYINC, described as income per household which is the product of real per capita income and the number of persons per residential connection. Four scenarios for price and income were used to derive four alternative forecasts of water consumption for El Paso, Abilene, Dallas, San Antonio, Corpus Christi and Houston. The assumptions for marginal price (MP) in the forecast period were 1) flat real prices throughout the forecast period 1989-1999, using the monthly values for 1988, the last year of historic data; and 2) annual growth in real prices of 4.1% (the average rate of increase for the 72 cities during 1980-1988). The assumptions for per capita income were 1) annual growth of 1.5% over the forecast period (the Texas Comptroller's current 20 year forecast rate of increase), and 2) zero growth, keeping income unchanged from the December 1988 level.

Persons per residential connection, and the commercial growth variable (COMPROXY) were held constant at their December 1988 levels. The number of dry days (DAYS) and the average monthly temperatures (TEMP) were forecast by projecting the average monthly values for the historic period.

Projections of Per Capita Water Consumption

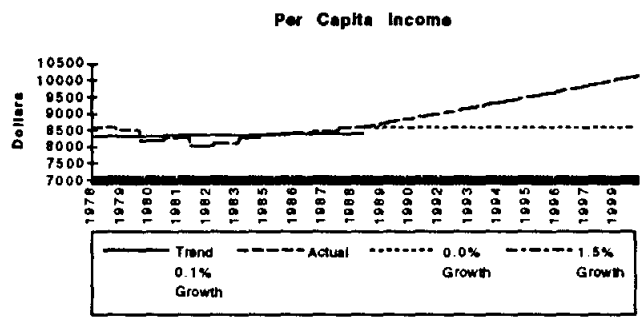
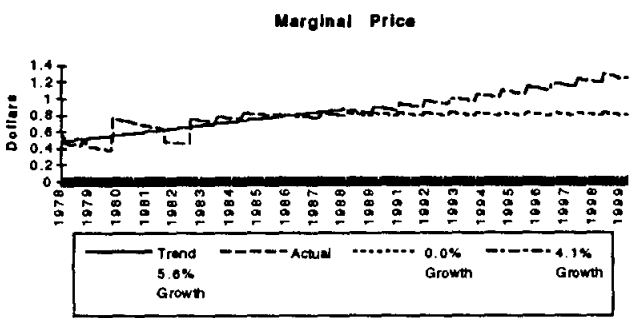
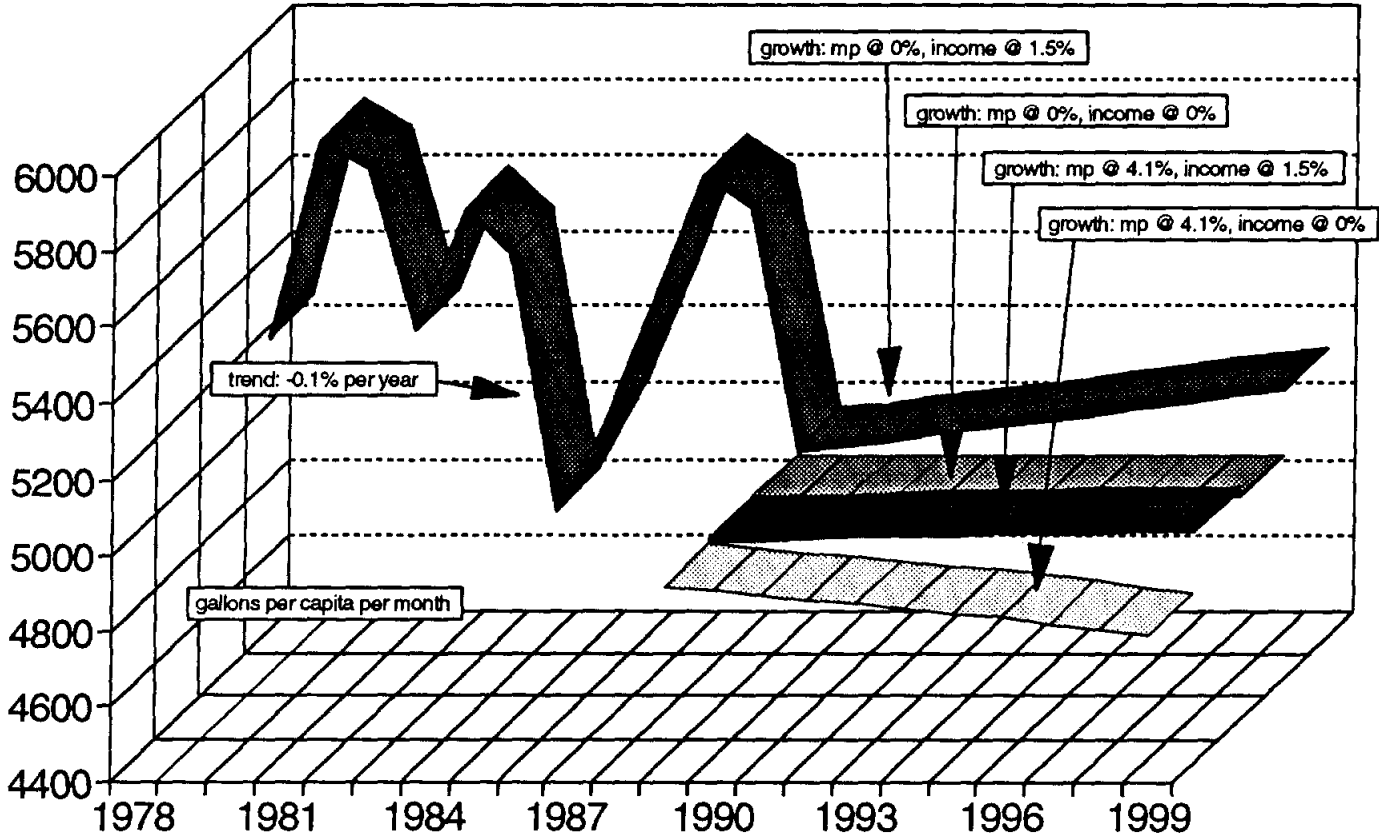
Forecasts of per capita water consumption for El Paso, Abilene, Dallas, San Antonio, Austin, Corpus and Houston are shown with projected assumptions of price, income and persons per residential connection in **Figures VI-1 and VI-7**. Historical data are also shown for perspective. The forecast results are summarized below by price/income scenario:

- a. Growth rates of 0% for price, 1.5% for personal income. Water consumption forecasts for all seven

Figure VI-1

EL PASO WATER CONSUMPTION

four scenarios for price, income

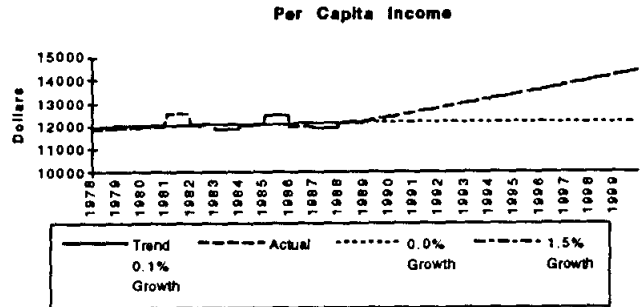
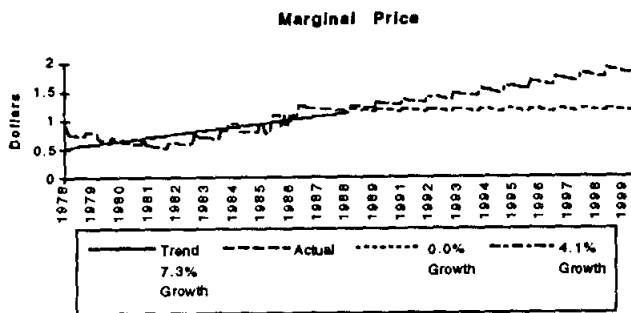
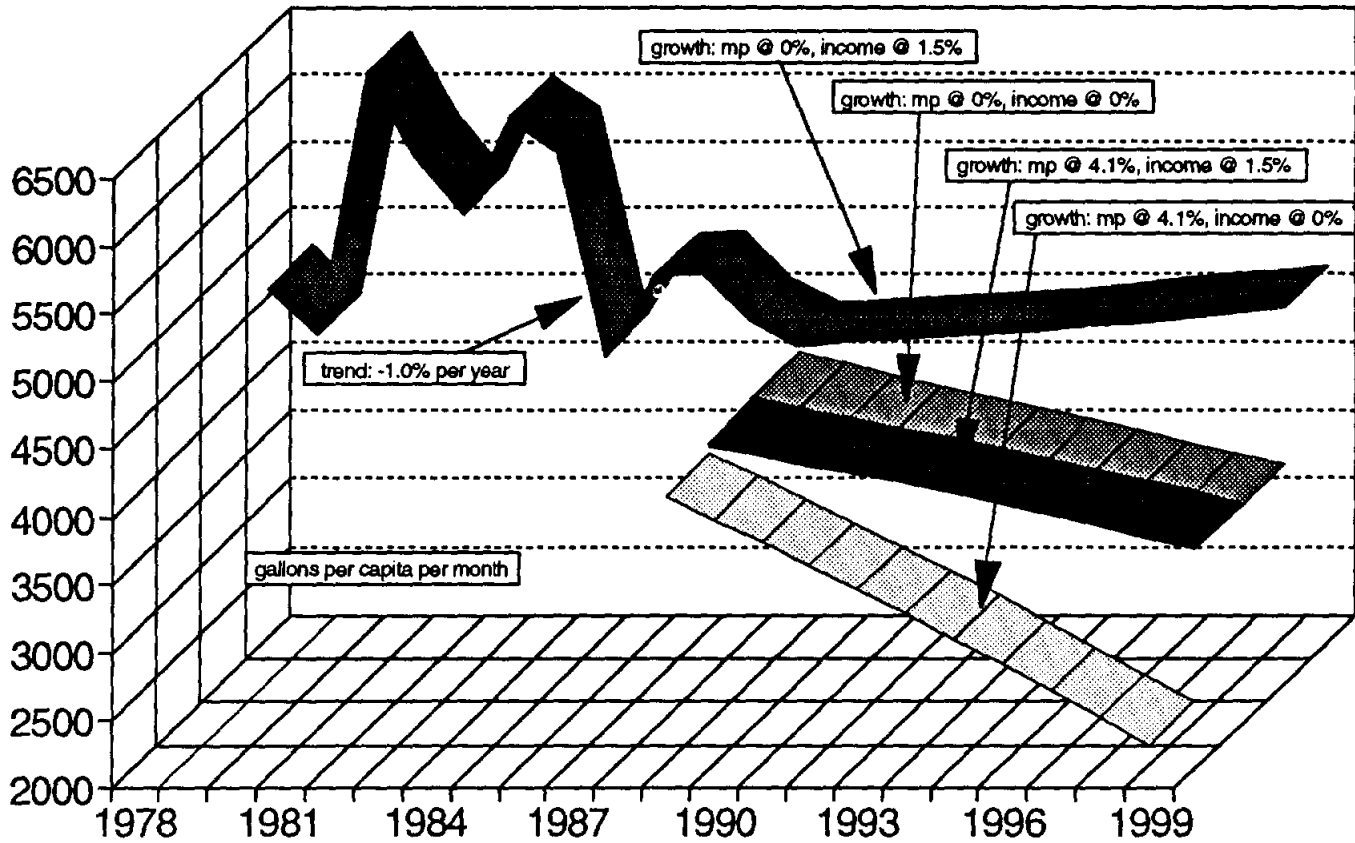


Persons Per Residential Connection held constant in projections

Figure VI-2

ABILENE WATER CONSUMPTION

four scenarios for price, income

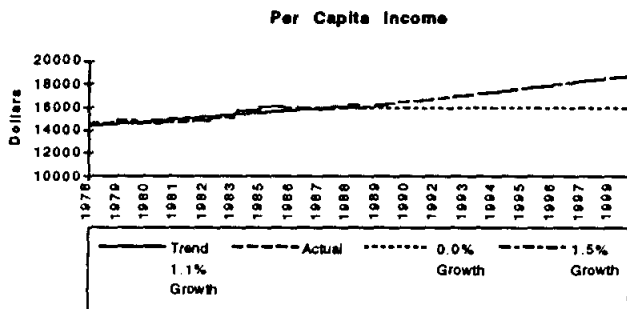
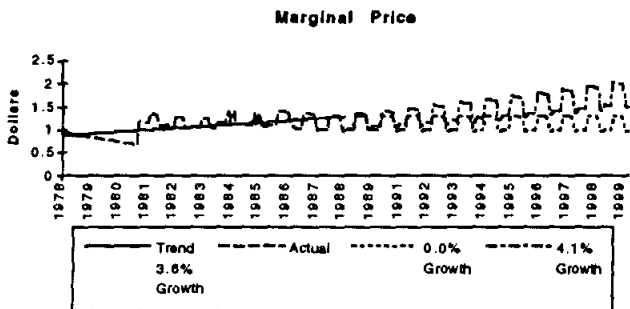
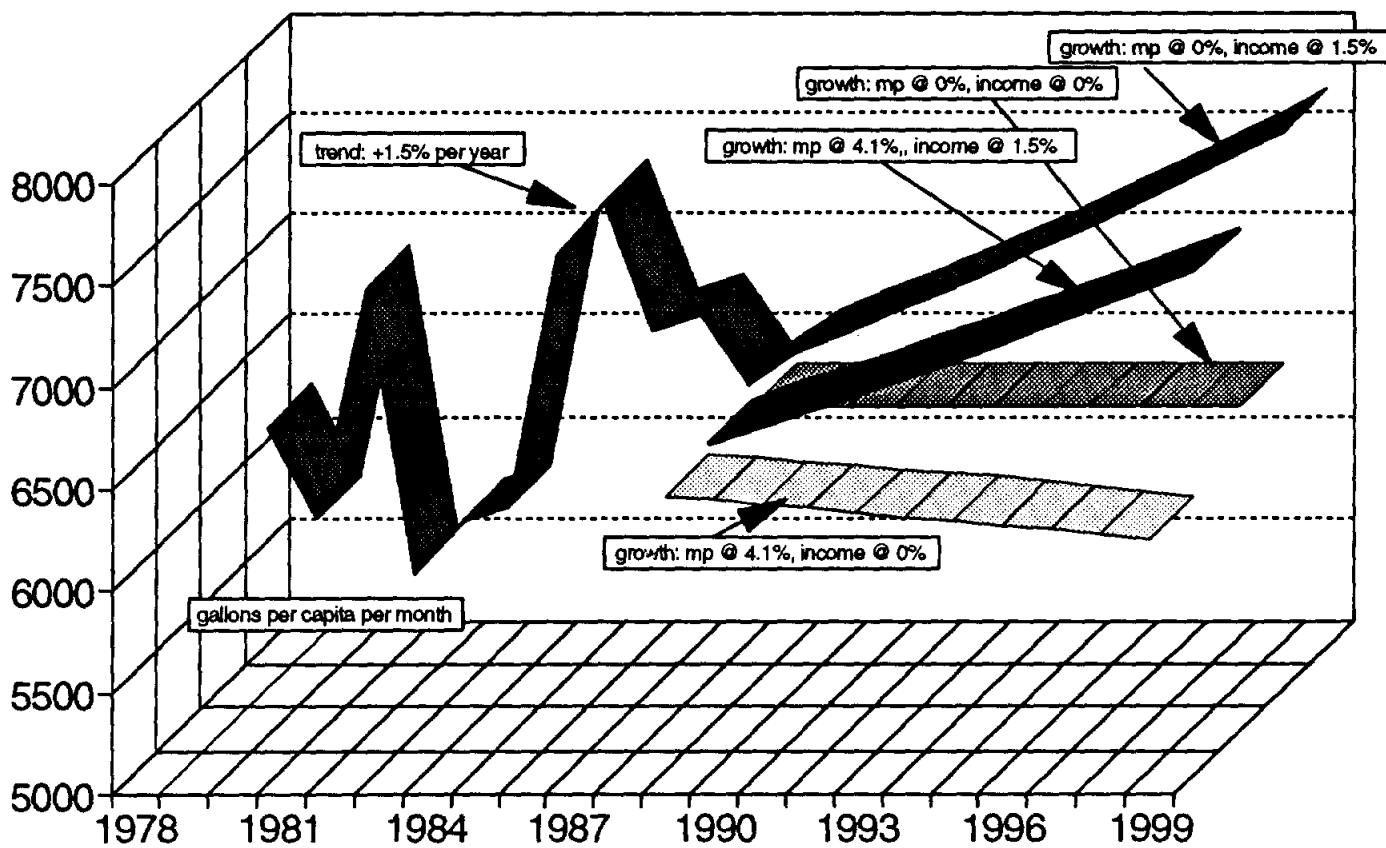


Persons Per Residential Connection held constant in projections

Figure VI-3

DALLAS WATER CONSUMPTION

four scenarios for price, income

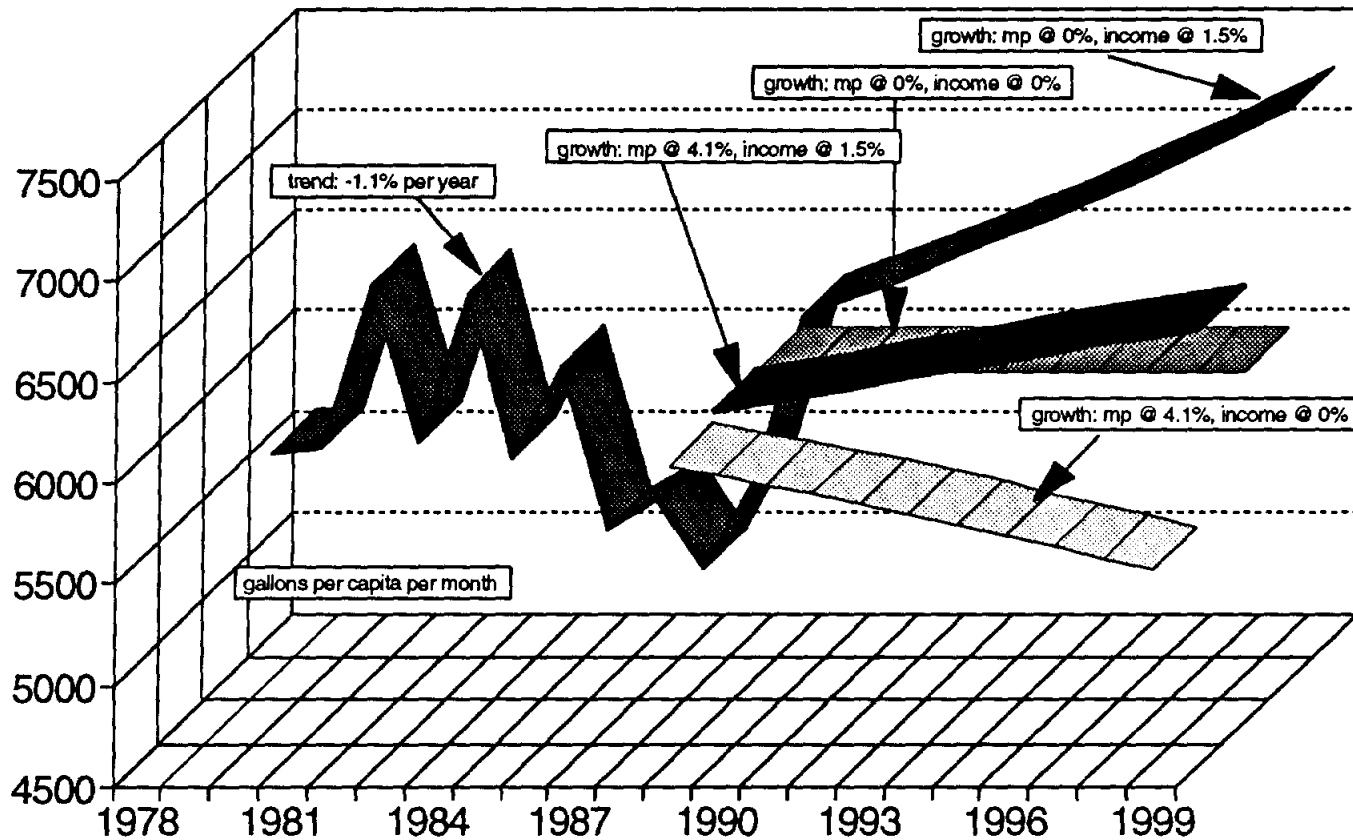


Persons Per Residential Connection held constant in projections

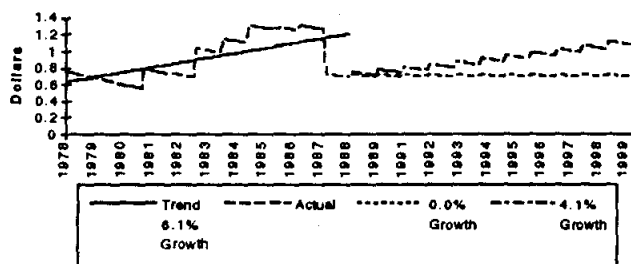
Figure VI-4

SAN ANTONIO WATER CONSUMPTION

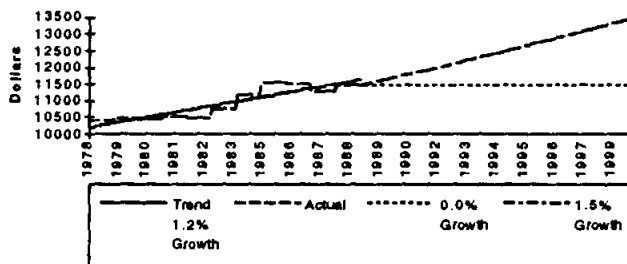
four scenarios for price, income



Marginal Price



Per Capita Income

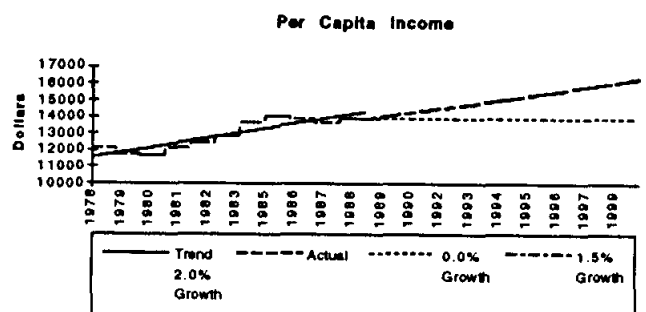
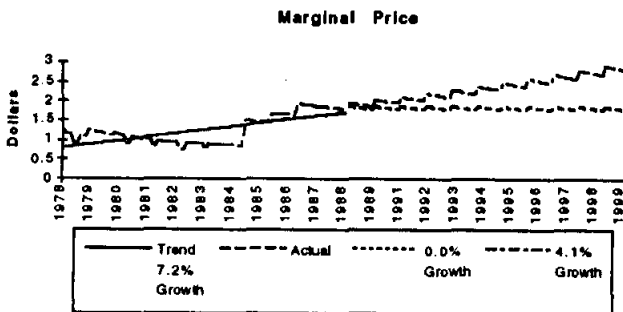
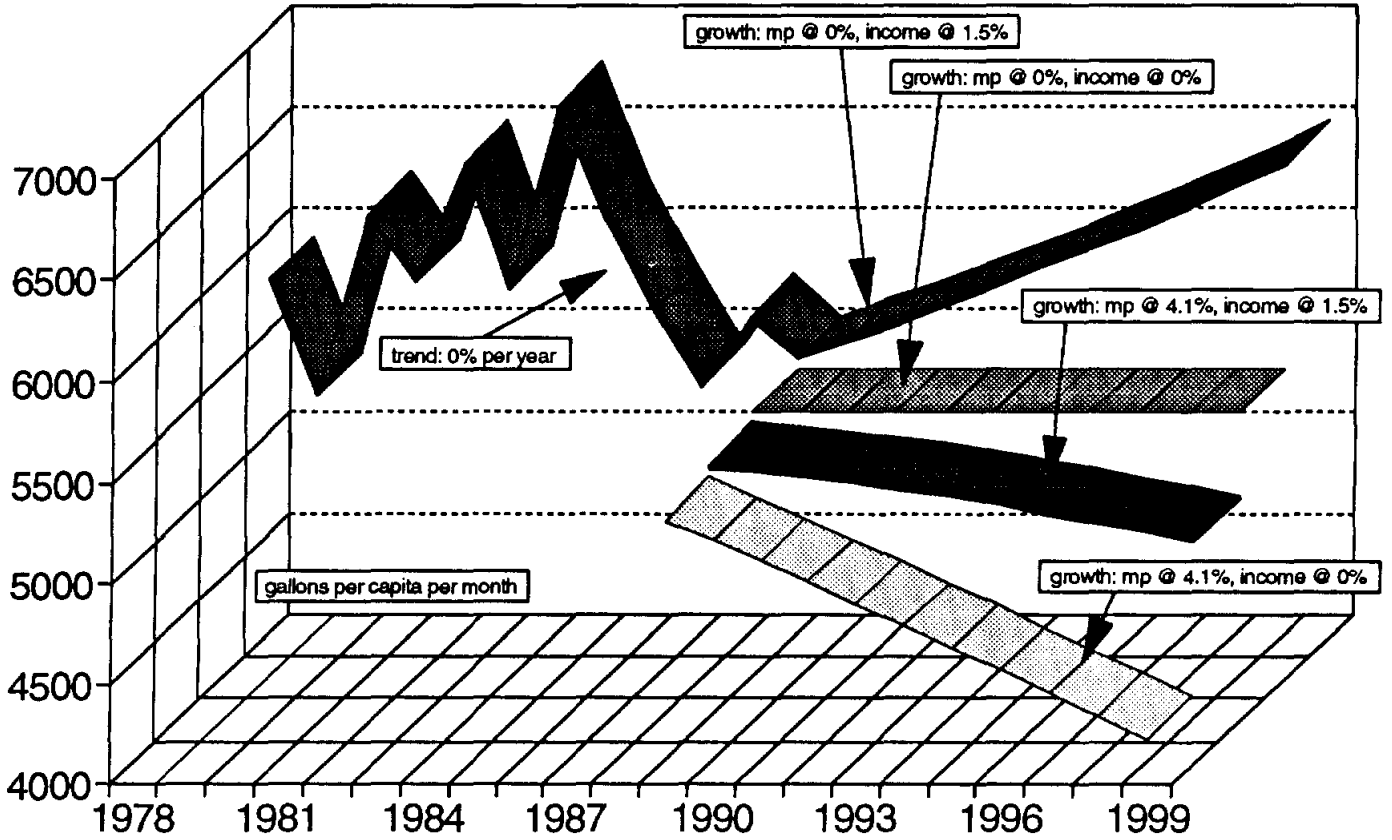


Persons Per Residential Connection held constant in projections

Figure VI-5

AUSTIN WATER CONSUMPTION

four scenarios for price, income

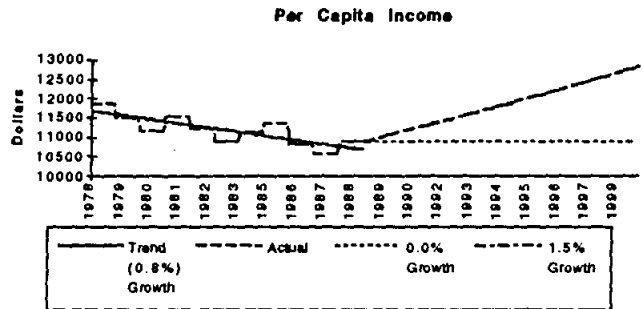
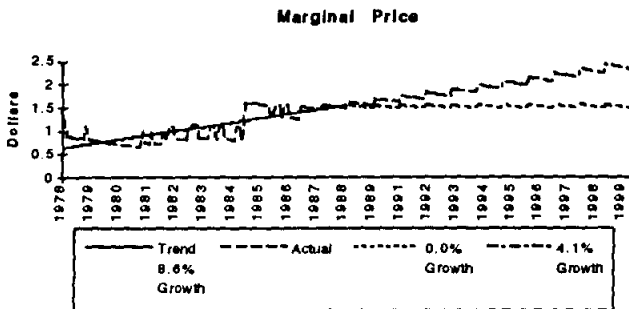
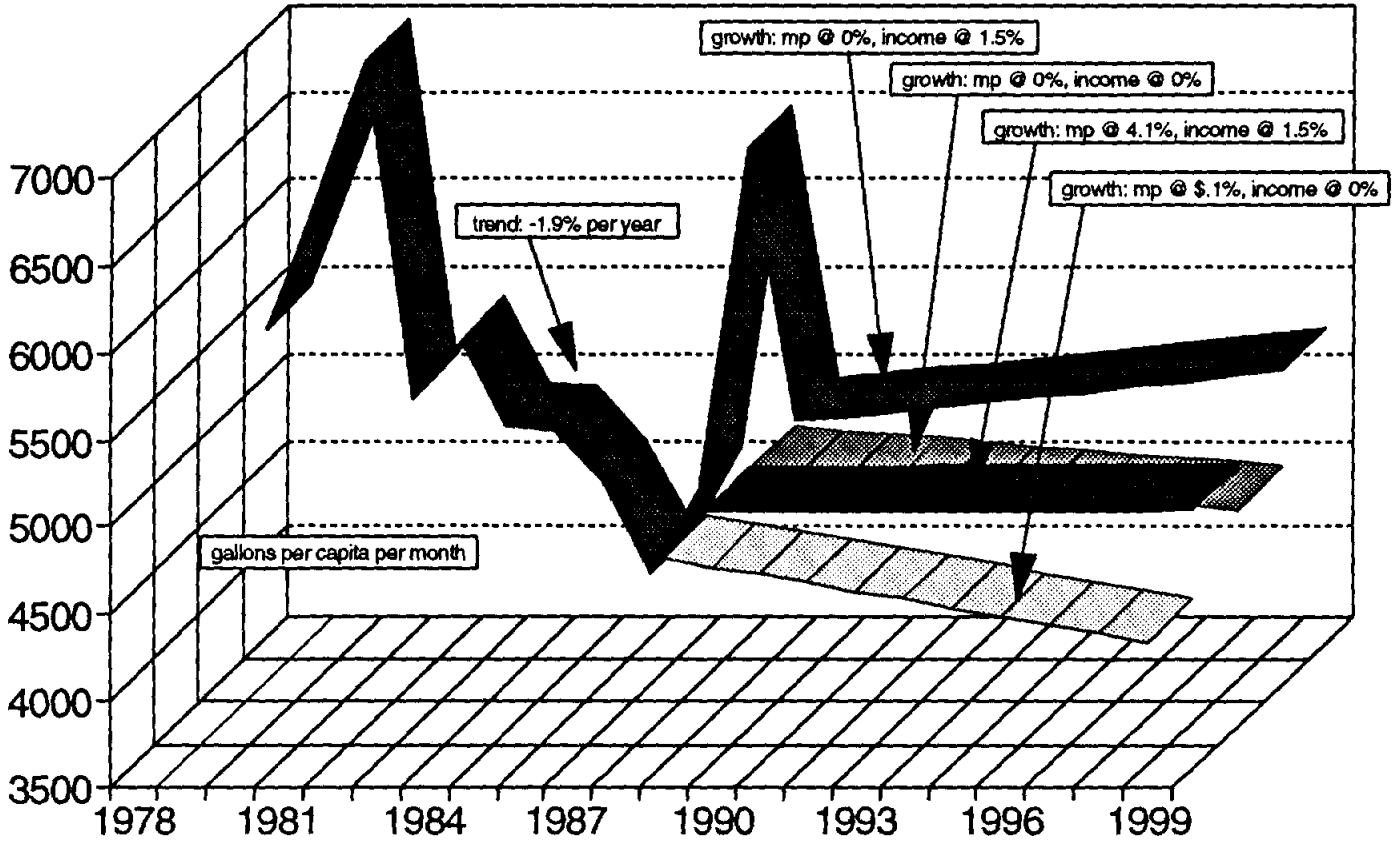


Persons Per Residential Connection held constant in projections

Figure VI-6

CORPUS WATER CONSUMPTION

four scenarios for price, income

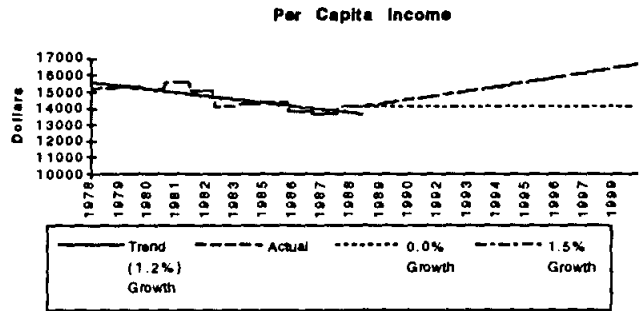
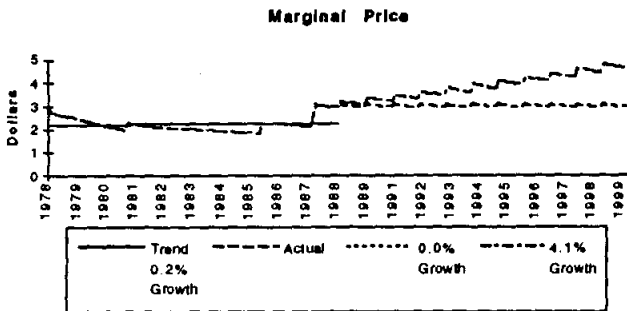
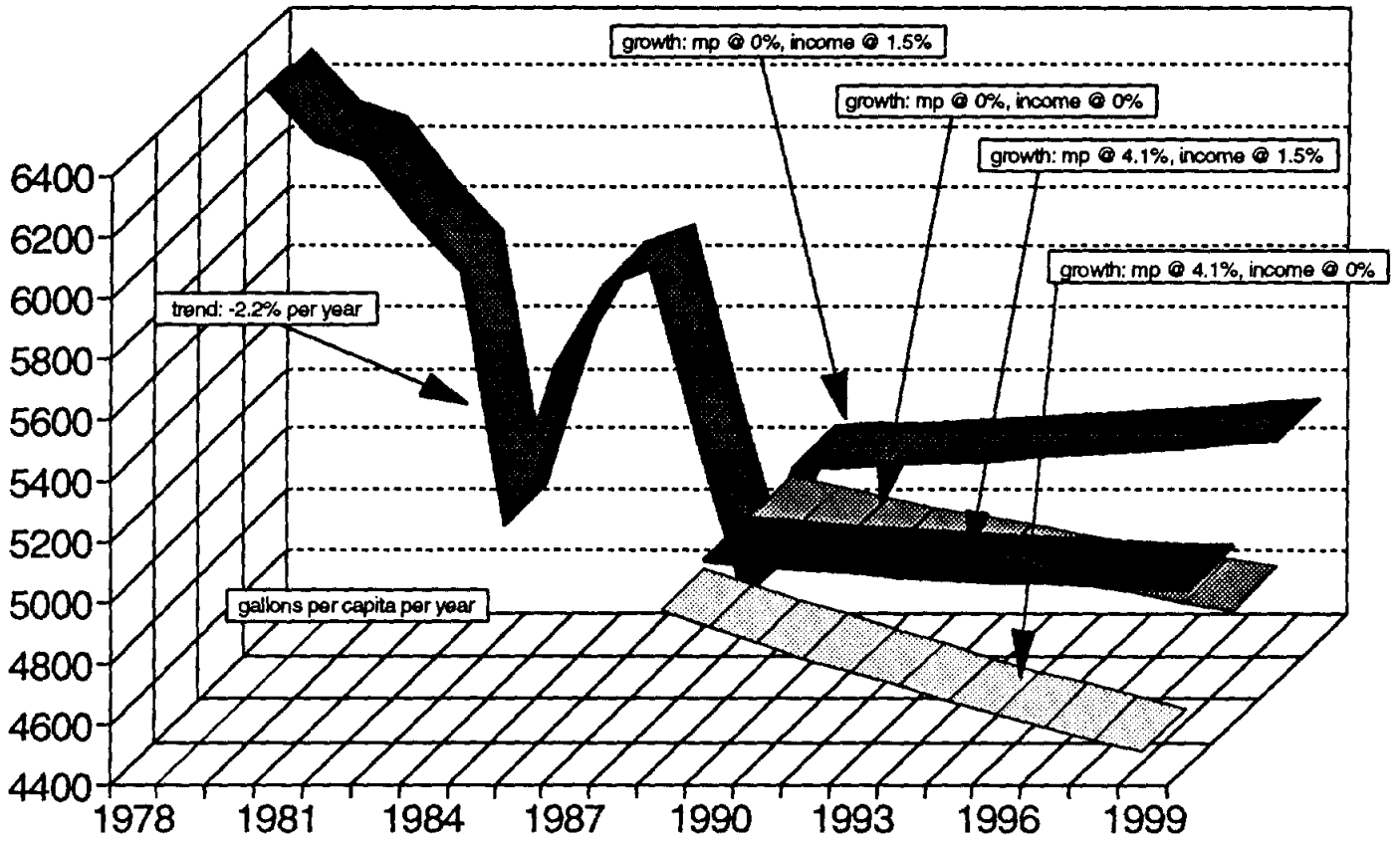


Persons Per Residential Connection held constant in projections

Figure VI-7

HOUSTON WATER CONSUMPTION

four scenarios for price, income



Persons Per Residential Connection held constant in projections

cities is projected to rise over the 1989-1999 period. Especially sharp increases are shown for Dallas, San Antonio and Austin relative to the other cities. The level of persons per residential connection for all these cities fell from 1978 to 1988. The arresting of that downward trend with a flat-growth assumption boosts the FAMLYINC variable for each city and therefore, the level of consumption. Persons per residential connection trended higher or remained virtually stable through 1988 for Houston, El Paso, Corpus and Abilene. The flat-growth assumption had a dampening effect, if any on income in the regional equations.

Note from **Table V-4** that the income elasticities of demand for water for the Dallas, San Antonio, Austin and Houston MSAs are higher than for the El Paso, Abilene and Corpus MSAs, and higher than most MSAs in the table. Income elasticity of demand measures the response of consumption to a one-percent change in income per residential connection. With other factors held constant, a 1.5% increase in income at San Antonio will yield an increase in water consumption of 1.4%. A 1.5% increase in income for El Paso yields only a 0.3% increase in water consumption, and **Figure VI-1** reflects the smaller response to income change.

- b. Growth rates of 0% for price and 0% for personal income. This scenario is unlikely over the period 1989-1999. This scenario produces forecasts for El Paso, Dallas, San Antonio and Austin of zero-growth in consumption because income and price are held constant, along with other predictive variables in the forecast period. Only for the cities of Abilene, Corpus and Houston, whose equations include the variable TIME, does consumption change over the forecast period. Consumption declines for Abilene, Corpus and Houston because TIME is inversely related to water consumption in their respective regional models.
- c. Growth rates of 4.1% for price and 1.5% for personal income. This scenario is very plausible. Under these assumptions of price and income growth, per capita water consumption would increase through the forecast period for El Paso, Dallas and San Antonio; would decrease for Abilene, Austin and Houston; and would remain virtually unchanged for Corpus.

The effect of the assumption of flat-growth in persons per residential connection is important again as it was in scenario a, but its importance is countered by the relative importance of price growth as seen by comparing price elasticities in **Table V-3** with

income elasticities in **Table V-4**.

Price elasticity of demand for water is a measure of the percent change in water consumption associated with a one-percent change in the price of water. Dallas consumption under scenario "a" shows an upward growth rate of **1.6%** annually in **Figure VI-3**, assuming 1.5% growth in income and 0% growth in price. With the assumption of 4.1% annual growth in price, consumption still shows upward growth over the forecast period. This persistent growth in consumption for Dallas is reflected in a relative insensitivity to water price changes combined with a relatively high income elasticity. The price elasticity for Dallas is quantified as -0.066. Of the MSAs representing the forecast cities, only the El Paso MSA has a lower price elasticity than that of Dallas (**Table V-3**). Price elasticity for Austin is nearly 4.5 times greater than for Dallas. The result of a 4.1% increase in price for Austin would result in declining consumption over the forecast period, other factors held constant.

- d. Growth rates of 4.1% for price and 0% for personal income. It is perhaps unlikely to have escalating real prices occurring with flat real personal income in the forecast period, but it is possible since that is basically the experience of the 1980s. If such a scenario were to happen, consumption would decline in all seven forecast cities.

Conservation Program Adjustments to Forecasts

Attempts to quantify conservation program effects using econometric methods applied in this analysis were not successful. Either the data are too weak, the effects not yet evident or the effects are not very important, independent of the other variables included in the models. One difficulty is the non-uniform definition of what constitutes a conservation program. The most important ambiguity is probably pricing. The analysis here reported includes marginal prices and, in many cases, utilities switched from flat or declining block rate structures to increasing block structures during the period of analysis, and classify such a change as a conservation program. If one accepts such a definitive, then the analysis in this report quantifies such a relationship.

The effects of mandatory government rules concerning appliance standards are not explicitly included in the current analysis. Projections of the effects of such mandates may be included for planning purposes by subtracting expected impacts from forecasts made using the

equations estimated in this study, but such a practice is an ad hoc method that is apt to overstate the case. That is, one would not expect per capita consumption reductions forthcoming from higher prices to be the same response that would ensue after water saving devices have been installed under government rules. The problem is that we cannot be sure how much reliance to place on the elasticity estimates in a market where behavior has been changed by government rule when the estimates of price response came for a market where such rules did not exist. The other difficulty is that municipal water is supplied by cost of service regulated utilities who may change rates in the future in a different time path under mandatory appliance standard than would be the case without them. The point is that price and price elasticities may be different with and without mandatory appliance standards so that accounting for the impacts of one cannot be considered independent of the other. This topic may need further research.

VII. SUMMARY AND FINDINGS

Summary

Eleven years of monthly consumption data for each of 72 cities in 28 MSAs of Texas were analyzed in this study in order to determine the underlying causes of declining per capita water use. Nine regional econometric models were estimated by grouping the 28 MSA sets of cities into homogeneous climatic and geographical groups. Each model allows an explanation of historical water consumption for each city in the group.

The set of six regional models all contain variables that we expect a priori to be important determinants of per capita water consumption. The equation forms and specific variables included differ among regions, and in some cases variables were ultimately dropped from the final equation because the estimated parameters were statistically insignificant; that is, we could not say with confidence that they had anything at all to do with consumption. In the end, however, the variables we believe should explain water consumption do in fact test significant and include (1) marginal price, (2) household income, (3) number of dry (low rainfall) days in the month, (4) temperature and (5) the concentration of commercial customers on the system. The fact that different forms and model specifications apply to different regions of the state also means that regional location and city size are important in explaining per capita municipal water consumption. The statistical properties of the models are all quite acceptable, and in fact are improved over many such results cited in the literature.

Price elasticity of demand estimates from the 72 cities in nine regions range from -0.042 to -0.543 while income elasticity estimates range from 0.031 to 1.267. These elasticities are well within the range of estimates obtained by others in the econometrics field. These elasticities allow simple calculation of the expected demand response to price and income changes.

Findings

The study of per capita municipal water demand during the 1978-1988 period leads to some interesting and important findings. The first important finding is that price, household income, concentration of commercial users, weather conditions, city size and location are all important variables in explaining historical water consumption and for forecasting future consumption. Seasonal variations are mostly explained by temperature and the lack of rainfall.

Long term trends are explained by household income, price and concentration of commercial users.

The general downward trend in per capita water consumption during 1978-1988 was the result of two sets of forces working at different parts of the time period, but themselves interrelated. The late 1970s and early 1980s brought rapid economic growth to Texas cities, resulting in at one and the same time, higher per capita incomes (exceeding the national average in 1982 for the first time ever) and explosive growth. Municipalities responded by constructing new facilities planned to catch up with growth and to meet a continued high growth in demand. By the mid-1980s growth had stopped and debt service requirements began to be realized, forcing utility rates to rise. Water supply, treatment and wastewater disposal costs also increased due to growing scarcity of supply and more stringent wastewater regulations. During this period of rising rates, many utilities switched from flat or declining block rate structures to increasing block structures, meaning that the marginal price of water rose above the average cost, a reversal of the historical relationship of the two prices. This sequence of events - rapid income and population growth - followed by stagnation and the lagged supply price response by cost of service based utilities meant that consumers were hit with stagnating incomes and rising marginal prices of water at the same time. The net result was a decline in per capita consumption rates.

If the above explanation of the past eleven years is correct, the question arises, "Will this downward trend in per capita consumption continue?" The analysis suggests that if the same forces of price, income, weather and persons per connection continue to determine consumption, per capita consumption is likely to continue declining in the foreseeable future.

There are two reasons why the trend, as influenced by the above factors, will continue. First, although Texas is now coming out of the longest recession it has had since WWII, the long term prospects for a per capita income growth near that of the post WWII era will be difficult to attain. Second, the overbuilding of utility capacity which occurred in the 1980s, leaving us with considerable excess capacity and high prices, should begin to abate soon, perhaps relieving the upward price pressure for a time, but in the long term prices are destined to rise rapidly. Most of the real price increases needed to retire the debt of the overbuilding have already been realized and real rates should begin to decline. The net result is that per capita consumption is likely to decline or stabilize in the long term.

The above conclusion will be reinforced by public policy driven by a number of interests ranging from public finance to environmental concerns. A case in point is an initiative to require certain water saving technologies to be installed by users, perhaps with the help of a public subsidy. Such mandates could further alter the consumption levels and trends of the future.

APPENDIX A

Statistical Output of Six Regional Models

Model: CENTRAL REGION: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|------|----------------|-------------|---------|--------|
| Model | 14 | 180.15052 | 12.86789 | 182.755 | 0.0001 |
| Error | 1305 | 91.88586 | 0.07041 | | |
| C Total | 1319 | 272.03638 | | | |
| Root MSE | | 0.26535 | R-square | 0.6622 | |
| Dep Mean | | 9.68666 | Adj R-sq | 0.6586 | |
| C.V. | | 2.73933 | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|-----------------------|-----------|
| INTERCEP | 1 | 8.070488 | 0.13094643 | 61.632 | 0.0000 |
| MP | 1 | -0.121359 | 0.03815518 | -3.181 | 0.0015 |
| FAMILYINC | 1 | 0.000028325 | 0.00000195 | 14.542 | 0.0001 |
| TEMP | 1 | 0.009232 | 0.00063210 | 14.606 | 0.0001 |
| DAYS | 1 | 0.010401 | 0.00326716 | 3.183 | 0.0015 |
| COMPROXY | 1 | 0.758233 | 0.24811419 | 3.056 | 0.0023 |
| BRY | 1 | -0.252587 | 0.03905116 | -6.468 | 0.0001 |
| BEL | 1 | -0.348467 | 0.04392496 | -7.933 | 0.0001 |
| WAC | 1 | 0.363213 | 0.03308491 | 10.978 | 0.0001 |
| WOO | 1 | 0.176476 | 0.03376189 | 5.227 | 0.0001 |
| COP | 1 | -0.385938 | 0.04336887 | -8.899 | 0.0001 |
| KIL | 1 | -0.531600 | 0.02927929 | -18.156 | 0.0001 |
| SHE | 1 | -0.383254 | 0.03369371 | -11.375 | 0.0001 |
| DEN | 1 | -0.352713 | 0.03511444 | -10.045 | 0.0001 |
| DSEAS | 1 | 0.179495 | 0.02517266 | 7.131 | 0.0001 |

Model: CENTRAL REGION: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|------|----------------|-------------|---------|--------|
| Model | 14 | 177.92388 | 12.70885 | 176.226 | 0.0001 |
| Error | 1305 | 94.11250 | 0.07212 | | |
| C Total | 1319 | 272.03638 | | | |
| Root MSE | | 0.26855 | R-square | 0.6540 | |
| Dep Mean | | 9.68666 | Adj R-sq | 0.6503 | |
| C.V. | | 2.77233 | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -3.419455 | 0.95563867 | -3.578 | 0.0004 |
| LOGMP | 1 | -0.160874 | 0.03960477 | -4.062 | 0.0001 |
| LOGFINC | 1 | 1.037026 | 0.08182628 | 12.674 | 0.0001 |
| LOGTEMP | 1 | 0.510024 | 0.03773387 | 13.516 | 0.0001 |
| LOGDAYS | 1 | 0.187562 | 0.07605561 | 2.466 | 0.0138 |
| LOGPROXY | 1 | 0.090974 | 0.02269853 | 4.008 | 0.0001 |
| BRY | 1 | -0.322024 | 0.04064294 | -7.923 | 0.0001 |
| BEL | 1 | -0.332745 | 0.03770657 | -8.825 | 0.0001 |
| WAC | 1 | 0.352985 | 0.03332756 | 10.591 | 0.0001 |
| WOO | 1 | 0.217403 | 0.03512248 | 6.190 | 0.0001 |
| COP | 1 | -0.424992 | 0.04810884 | -8.834 | 0.0001 |
| KIL | 1 | -0.539542 | 0.02973162 | -18.147 | 0.0001 |
| SHE | 1 | -0.435916 | 0.03466293 | -12.576 | 0.0001 |
| DEN | 1 | -0.401759 | 0.03656183 | -10.988 | 0.0001 |
| DSEAS | 1 | 0.222774 | 0.02422265 | 9.197 | 0.0001 |

Model: CENTRAL REGION: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-------------|----------------|--------------|---------|--------|
| Model | 14 | 85847139955 | 6131938568.2 | 157.729 | 0.0001 |
| Error | 1305 | 50733600381 | 38876322.131 | | |
| C Total | 1319 | 136580740336 | | | |
| Root MSE | 6235.08798 | R-square | 0.6285 | | |
| Dep Mean | 18057.72424 | Adj R-sq | 0.6246 | | |
| C.V. | 34.52865 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -21392 | 3076.9268493 | -6.952 | 0.0001 |
| MP | 1 | -2094.880831 | 896.55505433 | -2.337 | 0.0196 |
| FAMLYINC | 1 | 0.841070 | 0.04576768 | 18.377 | 0.0001 |
| TEMP | 1 | 158.520521 | 14.85293142 | 10.673 | 0.0001 |
| DAYS | 1 | 188.335911 | 76.77035019 | 2.453 | 0.0143 |
| COMPROXY | 1 | 325.211345 | 5830.0878418 | 0.056 | 0.9555 |
| BRY | 1 | -6260.006893 | 917.60844340 | -6.822 | 0.0001 |
| BEL | 1 | -4441.321867 | 1032.1311749 | -4.303 | 0.0001 |
| WAC | 1 | 5503.321598 | 777.41604396 | 7.079 | 0.0001 |
| WOO | 1 | 4455.657161 | 793.32327048 | 5.616 | 0.0001 |
| COP | 1 | -5169.700281 | 1019.0643732 | -5.073 | 0.0001 |
| KIL | 1 | -12312 | 687.99309407 | -17.896 | 0.0001 |
| SHE | 1 | -7779.144144 | 791.72125551 | -9.826 | 0.0001 |
| DEN | 1 | -6812.752594 | 825.10503953 | -8.257 | 0.0001 |
| DSEAS | 1 | 4504.858248 | 591.49713271 | 7.616 | 0.0001 |

CENTRAL REGION

CORRELATION ANALYSIS

5 'VAR' Variables: MP FAMLYINC TEMP DAYS COMPROXY

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|------|----------|----------|-----------|
| MP | 1320 | 1.11217 | 0.23656 | 1468 |
| FAMLYINC | 1320 | 36966 | 6201 | 48794707 |
| TEMP | 1320 | 65.69037 | 14.48127 | 86711 |
| DAYS | 1320 | 27.12879 | 2.34320 | 35810 |
| COMPROXY | 1320 | 0.09497 | 0.05994 | 125.35589 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| MP | 0.33841 | 2.07254 |
| FAMLYINC | 25042 | 54936 |
| TEMP | 30.87000 | 89.23000 |
| DAYS | 5.00000 | 31.00000 |
| COMPROXY | 0.01000 | 0.30000 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0 / N = 1320

| | MP | FAMLYINC | TEMP | DAYS | COMPROXY |
|----------|--------------------|-------------------|--------------------|--------------------|--------------------|
| MP | 1.00000 0.0 | 0.23771 0.0001 | -0.01572 0.5683 | 0.00671 0.8076 | -0.20257 0.0001 |
| FAMLYINC | 0.23771 0.0001 | 1.00000 0.0 | 0.05296 0.0544 | 0.06463 0.0189 | 0.22793 0.0001 |
| TEMP | -0.01572 0.5683 | 0.05296 0.0544 | 1.00000 0.0 | 0.07000 0.0110 | 0.02417 0.3803 |
| DAYS | 0.00671 0.8076 | 0.06463 0.0189 | 0.07000 0.0110 | 1.00000 0.0 | -0.02355 0.3926 |
| COMPROXY | -0.20257 0.0001 | 0.22793 0.0001 | 0.02417 0.3803 | -0.02355 0.3926 | 1.00000 0.0 |

Model: EAST REGION: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-------------|----------------|--------------|---------|--------|
| Model | 14 | 52912708524 | 3779479180.3 | 325.748 | 0.0000 |
| Error | 1701 | 19735785576 | 11602460.656 | | |
| C Total | 1715 | 72648494100 | | | |
| Root MSE | 3406.23849 | R-square | 0.7283 | | |
| Dep Mean | 16260.43531 | Adj R-sq | 0.7261 | | |
| C.V. | 20.94802 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -1579.937997 | 1124.6315610 | -1.405 | 0.1602 |
| TIME | 1 | -0.364233 | 0.07521000 | -4.843 | 0.0001 |
| MP | 1 | -1150.623861 | 253.97766279 | -4.530 | 0.0001 |
| FAMLYINC | 1 | 0.129296 | 0.01038690 | 12.448 | 0.0001 |
| TEMP | 1 | 109.997310 | 7.55527808 | 14.559 | 0.0001 |
| DAYS | 1 | 112.519242 | 28.58107062 | 3.937 | 0.0001 |
| COMPROXY | 1 | 20658 | 1650.2929431 | 12.518 | 0.0001 |
| D25 | 1 | 5877.016682 | 354.01056669 | 16.601 | 0.0001 |
| D4 | 1 | 2154.868306 | 274.49483884 | 7.850 | 0.0001 |
| TXK | 1 | 8736.179747 | 428.44861753 | 20.390 | 0.0001 |
| CON | 1 | 9358.399530 | 395.11488863 | 23.685 | 0.0001 |
| HOU | 1 | 12702 | 508.29071316 | 24.991 | 0.0001 |
| BAY | 1 | 1408.586397 | 370.94263328 | 3.797 | 0.0002 |
| D16 | 1 | 4286.580079 | 277.69934681 | 15.436 | 0.0001 |
| DSEAS | 1 | 2597.416524 | 269.67593575 | 9.632 | 0.0001 |

Model: EAST REGION: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|------|----------------|-------------|---------|--------|
| Model | 14 | 189.02420 | 13.50173 | 338.696 | 0.0000 |
| Error | 1701 | 67.80846 | 0.03986 | | |
| C Total | 1715 | 256.83266 | | | |
| Root MSE | | 0.19966 | R-square | 0.7360 | |
| Dep Mean | | 9.62164 | Adj R-sq | 0.7338 | |
| C.V. | | 2.07511 | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 8.316178 | 0.06592118 | 126.153 | 0.0000 |
| TIME | 1 | -0.000020700 | 0.00000441 | -4.696 | 0.0001 |
| MP | 1 | -0.054906 | 0.01488710 | -3.688 | 0.0002 |
| FAMLYINC | 1 | 0.000011725 | 0.00000061 | 19.258 | 0.0001 |
| TEMP | 1 | 0.007283 | 0.00044286 | 16.445 | 0.0001 |
| DAYS | 1 | 0.005839 | 0.00167530 | 3.485 | 0.0005 |
| COMPROXY | 1 | 0.996068 | 0.09673325 | 10.297 | 0.0001 |
| D25 | 1 | 0.394668 | 0.02075061 | 19.020 | 0.0001 |
| D4 | 1 | 0.225046 | 0.01608974 | 13.987 | 0.0001 |
| TXK | 1 | 0.671280 | 0.02511386 | 26.729 | 0.0001 |
| CON | 1 | 0.556300 | 0.02315998 | 24.020 | 0.0001 |
| HOU | 1 | 0.604033 | 0.02979387 | 20.274 | 0.0001 |
| BAY | 1 | 0.132475 | 0.02174310 | 6.093 | 0.0001 |
| D16 | 1 | 0.365347 | 0.01627757 | 22.445 | 0.0001 |
| DSEAS | 1 | 0.127176 | 0.01580727 | 8.045 | 0.0001 |

Model: EAST REGION: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|------|----------------|-------------|---------|--------|
| Model | 14 | 189.06908 | 13.50493 | 338.025 | 0.0000 |
| Error | 1692 | 67.59957 | 0.03995 | | |
| C Total | 1706 | 256.66865 | | | |
| Root MSE | | 0.19988 | R-square | 0.7366 | |
| Dep Mean | | 9.62178 | Adj R-sq | 0.7344 | |
| C.V. | | 2.07738 | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 3.132205 | 0.47267635 | 6.627 | 0.0001 |
| LOGTIME | 1 | -0.211328 | 0.03766911 | -5.610 | 0.0001 |
| LOGMP | 1 | -0.090163 | 0.02059166 | -4.379 | 0.0001 |
| LOGFINC | 1 | 0.533279 | 0.02595968 | 20.543 | 0.0001 |
| LOGTEMP | 1 | 0.419240 | 0.02636114 | 15.904 | 0.0001 |
| LOGPROXY | 1 | 0.069177 | 0.00941798 | 7.345 | 0.0001 |
| LOGDAYS | 1 | 0.266285 | 0.05493009 | 4.848 | 0.0001 |
| D25 | 1 | 0.377579 | 0.02088575 | 18.078 | 0.0001 |
| D4 | 1 | 0.181877 | 0.01634712 | 11.126 | 0.0001 |
| TXK | 1 | 0.771895 | 0.02829364 | 27.282 | 0.0001 |
| CON | 1 | 0.564478 | 0.02305658 | 24.482 | 0.0001 |
| HOU | 1 | 0.613263 | 0.02761729 | 22.206 | 0.0001 |
| BAY | 1 | 0.107874 | 0.02177453 | 4.954 | 0.0001 |
| D16 | 1 | 0.371030 | 0.01575022 | 23.557 | 0.0001 |
| DSEAS | 1 | 0.145817 | 0.01527212 | 9.548 | 0.0001 |

EAST REGION

CORRELATION ANALYSIS

6 'VAR' Variables: LOGTIME LOGMP LOGFINC LOGTEMP LOGDAYS LOGPROXY

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|------|----------|---------|-----------|
| LOGTIME | 1716 | 0.25911 | 0.31671 | 444.63488 |
| LOGMP | 1716 | 0.25911 | 0.31671 | 444.63488 |
| LOGFINC | 1716 | 10.60316 | 0.35252 | 18195 |
| LOGTEMP | 1716 | 4.16216 | 0.22288 | 7142 |
| LOGDAYS | 1707 | 3.26597 | 0.08990 | 5575 |
| LOGPROXY | 1716 | -2.45621 | 0.63691 | -4215 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| LOGTIME | -0.54000 | 1.13051 |
| LOGMP | -0.54000 | 1.13051 |
| LOGFINC | 9.58014 | 11.30459 |
| LOGTEMP | 3.45774 | 4.46740 |
| LOGDAYS | 2.89037 | 3.43399 |
| LOGPROXY | -3.90197 | -0.98450 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0
/ Number of Observations

| | LOGTIME | LOGMP | LOGFINC | LOGTEMP | LOGDAYS | LOGPROXY |
|----------|----------------------------|----------------------------|---------------------------|----------------------------|---------------------------|----------------------------|
| LOGTIME | 1.00000 0.0 1716 | 1.00000 0.0 1716 | 0.08991 0.0002 1716 | -0.02804 0.2456 1716 | 0.02123 0.3808 1707 | -0.26653 0.0001 1716 |
| LOGMP | 1.00000 0.0 1716 | 1.00000 0.0 1716 | 0.08991 0.0002 1716 | -0.02804 0.2456 1716 | 0.02123 0.3808 1707 | -0.26653 0.0001 1716 |
| LOGFINC | 0.08991 0.0002 1716 | 0.08991 0.0002 1716 | 1.00000 0.0 1716 | 0.16801 0.0001 1716 | 0.04395 0.0695 1707 | 0.22642 0.0001 1716 |
| LOGTEMP | -0.02804 0.2456 1716 | -0.02804 0.2456 1716 | 0.16801 0.0001 1716 | 1.00000 0.0 1716 | 0.12853 0.0001 1707 | 0.05454 0.0239 1716 |
| LOGDAYS | 0.02123 0.3808 1707 | 0.02123 0.3808 1707 | 0.04395 0.0695 1707 | 0.12853 0.0001 1707 | 1.00000 0.0 1707 | 0.00125 0.9589 1707 |
| LOGPROXY | -0.26653 0.0001 1716 | -0.26653 0.0001 1716 | 0.22642 0.0001 1716 | 0.05454 0.0239 1716 | 0.00125 0.9589 1707 | 1.00000 0.0 1716 |

Model: I-35 REGION: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-------------|----------------|--------------|---------|--------|
| Model | 7 | 15232634292 | 2176090613.1 | 245.938 | 0.0001 |
| Error | 388 | 3433068803.4 | 8848115.4727 | | |
| C Total | 395 | 18665703095 | | | |
| Root MSE | 2974.57820 | R-square | 0.8161 | | |
| Dep Mean | 21774.67172 | Adj R-sq | 0.8128 | | |
| C.V. | 13.66073 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -20263 | 3215.9384331 | -6.301 | 0.0001 |
| MP | 1 | -5563.864316 | 620.15135631 | -8.972 | 0.0001 |
| FAMILYINC | 1 | 0.344895 | 0.05434269 | 6.347 | 0.0001 |
| TEMP | 1 | 227.528724 | 14.56079069 | 15.626 | 0.0001 |
| DAYS | 1 | 329.728042 | 71.40199608 | 4.618 | 0.0001 |
| D3 | 1 | 7392.749386 | 469.20350222 | 15.756 | 0.0001 |
| D22 | 1 | 8301.879576 | 471.03971698 | 17.625 | 0.0001 |
| DSEAS | 1 | 5449.013913 | 493.00895467 | 11.053 | 0.0001 |

Model: I-35 REGION: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-----|----------------|-------------|---------|--------|
| Model | 7 | 30.73045 | 4.39006 | 332.095 | 0.0001 |
| Error | 388 | 5.12909 | 0.01322 | | |
| C Total | 395 | 35.85954 | | | |
| Root MSE | | 0.11498 | R-square | 0.8570 | |
| Dep Mean | | 9.94242 | Adj R-sq | 0.8544 | |
| C.V. | | 1.15641 | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 7.964865 | 0.12430443 | 64.075 | 0.0001 |
| MP | 1 | -0.244931 | 0.02397047 | -10.218 | 0.0001 |
| FAMLYINC | 1 | 0.000018551 | 0.00000210 | 8.832 | 0.0001 |
| TEMP | 1 | 0.010514 | 0.00056281 | 18.682 | 0.0001 |
| DAYS | 1 | 0.011598 | 0.00275987 | 4.202 | 0.0001 |
| D3 | 1 | 0.352806 | 0.01813594 | 19.453 | 0.0001 |
| D22 | 1 | 0.388082 | 0.01820692 | 21.315 | 0.0001 |
| DSEAS | 1 | 0.187767 | 0.01905609 | 9.853 | 0.0001 |

Model: I-35 REGION: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 7 | 30.38941 | 4.34134 | 307.935 | 0.0001 |
| Error | 388 | 5.47013 | 0.01410 | | |
| C Total | 395 | 35.85954 | | | |
| Root MSE | 0.11874 | R-square | 0.8475 | | |
| Dep Mean | 9.94242 | Adj R-sq | 0.8447 | | |
| C.V. | 1.19424 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -3.620332 | 1.16299161 | -3.113 | 0.0020 |
| LOGMP | 1 | -0.261124 | 0.02689529 | -9.709 | 0.0001 |
| LOGFINC | 1 | 0.894294 | 0.10477027 | 8.536 | 0.0001 |
| LOGTEMP | 1 | 0.640729 | 0.03687404 | 17.376 | 0.0001 |
| LOGDAYS | 1 | 0.279285 | 0.07438467 | 3.755 | 0.0002 |
| D3 | 1 | 0.345736 | 0.01894379 | 18.251 | 0.0001 |
| D22 | 1 | 0.382372 | 0.01910090 | 20.019 | 0.0001 |
| DSEAS | 1 | 0.215919 | 0.01905356 | 11.332 | 0.0001 |

CORRELATION ANALYSIS

4 'VAR' Variables: MP FAMLYINC TEMP DAYS

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|-----|----------|----------|-----------|
| MP | 396 | 1.01047 | 0.32806 | 400.14477 |
| FAMLYINC | 432 | 48962 | 4015 | 21151478 |
| TEMP | 432 | 68.92685 | 12.65283 | 29776 |
| DAYS | 432 | 27.35880 | 2.12454 | 11819 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| MP | 0.55944 | 1.94192 |
| FAMLYINC | 41498 | 57675 |
| TEMP | 40.39000 | 88.11000 |
| DAYS | 19.00000 | 31.00000 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0
/ Number of Observations

| | MP | FAMLYINC | TEMP | DAYS |
|----------|---------------------------|---------------------------|---------------------------|---------------------------|
| MP | 1.00000 0.0 396 | 0.51068 0.0001 396 | -0.05152 0.3065 396 | 0.02505 0.6192 396 |
| FAMLYINC | 0.51068 0.0001 396 | 1.00000 0.0 432 | -0.04084 0.3972 432 | 0.05952 0.2170 432 |
| TEMP | -0.05152 0.3065 396 | -0.04084 0.3972 432 | 1.00000 0.0 432 | -0.02557 0.5961 432 |
| DAYS | 0.02505 0.6192 396 | 0.05952 0.2170 432 | -0.02557 0.5961 432 | 1.00000 0.0 432 |

Model: METROPLEX REGION: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-------------|----------------|--------------|---------|--------|
| Model | 7 | 52822846616 | 7546120945.2 | 433.794 | 0.0001 |
| Error | 916 | 15934390605 | 17395622.931 | | |
| C Total | 923 | 68757237221 | | | |
| Root MSE | 4170.80603 | R-square | 0.7683 | | |
| Dep Mean | 21575.16126 | Adj R-sq | 0.7665 | | |
| C.V. | 19.33152 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -33339 | 2581.8729505 | -12.913 | 0.0001 |
| MP | 1 | -959.272337 | 413.05544855 | -2.322 | 0.0204 |
| FAMILYINC | 1 | 0.336998 | 0.03246403 | 10.381 | 0.0001 |
| TEMP | 1 | 281.819178 | 11.50629529 | 24.493 | 0.0001 |
| DAYS | 1 | 506.864726 | 65.72032717 | 7.712 | 0.0001 |
| COMPROXY | 1 | 39487 | 5795.5374962 | 6.813 | 0.0001 |
| D11 | 1 | 2539.022545 | 370.14284720 | 6.860 | 0.0001 |
| DSEAS | 1 | 7058.573488 | 496.04309413 | 14.230 | 0.0001 |

Model: METROPLEX REGION: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 7 | 103.43517 | 14.77645 | 486.338 | 0.0001 |
| Error | 916 | 27.83092 | 0.03038 | | |
| C Total | 923 | 131.26608 | | | |
| Root MSE | 0.17431 | R-square | 0.7880 | | |
| Dep Mean | 9.90660 | Adj R-sq | 0.7864 | | |
| C.V. | 1.75951 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 7.315527 | 0.10790235 | 67.798 | 0.0000 |
| MP | 1 | -0.053469 | 0.01726253 | -3.097 | 0.0020 |
| FAMILYINC | 1 | 0.000014511 | 0.00000136 | 10.695 | 0.0001 |
| TEMP | 1 | 0.014284 | 0.00048087 | 29.703 | 0.0001 |
| DAYS | 1 | 0.024448 | 0.00274660 | 8.901 | 0.0001 |
| COMPROXY | 1 | 2.137456 | 0.24220871 | 8.825 | 0.0001 |
| D11 | 1 | 0.142162 | 0.01546911 | 9.190 | 0.0001 |
| DSEAS | 1 | 0.201805 | 0.02073077 | 9.735 | 0.0001 |

Model: METROPLEX REGION: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 7 | 99.60714 | 14.22959 | 411.710 | 0.0001 |
| Error | 916 | 31.65895 | 0.03456 | | |
| C Total | 923 | 131.26608 | | | |
| Root MSE | 0.18591 | R-square | 0.7588 | | |
| Dep Mean | 9.90660 | Adj R-sq | 0.7570 | | |
| C.V. | 1.87662 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|-----------------------|-----------|
| INTERCEP | 1 | -5.279767 | 0.89394871 | -5.906 | 0.0001 |
| LOGMP | 1 | -0.073329 | 0.02324518 | -3.155 | 0.0017 |
| LOGFINC | 1 | 0.929430 | 0.07492974 | 12.404 | 0.0001 |
| LOGTEMP | 1 | 0.792572 | 0.02995773 | 26.456 | 0.0001 |
| LOGDAYS | 1 | 0.595566 | 0.07601118 | 7.835 | 0.0001 |
| LOGPROXY | 1 | 0.105710 | 0.01459278 | 7.244 | 0.0001 |
| D11 | 1 | 0.128456 | 0.01753912 | 7.324 | 0.0001 |
| DSEAS | 1 | 0.270423 | 0.02097728 | 12.891 | 0.0001 |

METROPLEX REGION

CORRELATION ANALYSIS

5 'VAR' Variables: MP FAMLYINC TEMP DAYS COMPROXY

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|-----|----------|----------|----------|
| MP | 924 | 1.20908 | 0.33655 | 1117 |
| FAMLYINC | 924 | 55038 | 7975 | 50855321 |
| TEMP | 924 | 65.27315 | 15.11004 | 60312 |
| DAYS | 924 | 27.46861 | 2.26617 | 25381 |
| COMPROXY | 924 | 0.08372 | 0.04056 | 77.35477 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| MP | 0.64000 | 2.34005 |
| FAMLYINC | 39711 | 76343 |
| TEMP | 32.58500 | 90.90000 |
| DAYS | 15.00000 | 31.00000 |
| COMPROXY | 0.01000 | 0.20000 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0 / N = 924

| | MP | FAMLYINC | TEMP | DAYS | COMPROXY |
|----------|--------------------|--------------------|--------------------|--------------------|--------------------|
| MP | 1.00000 0.0 | 0.01219 0.7114 | 0.02425 0.4616 | -0.04852 0.1405 | 0.10195 0.0019 |
| FAMLYINC | 0.01219 0.7114 | 1.00000 0.0 | -0.00263 0.9365 | -0.00092 0.9777 | 0.76307 0.0001 |
| TEMP | 0.02425 0.4616 | -0.00263 0.9365 | 1.00000 0.0 | 0.09801 0.0029 | 0.01196 0.7165 |
| DAYS | -0.04852 0.1405 | -0.00092 0.9777 | 0.09801 0.0029 | 1.00000 0.0 | -0.03465 0.2927 |
| COMPROXY | 0.10195 0.0019 | 0.76307 0.0001 | 0.01196 0.7165 | -0.03465 0.2927 | 1.00000 0.0 |

Model: METROPLEX SURBURBAN: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-------------|----------------|--------------|---------|--------|
| Model | 8 | 46988787930 | 5873598491.3 | 411.303 | 0.0000 |
| Error | 1443 | 20606694942 | 14280453.875 | | |
| C Total | 1451 | 67595482872 | | | |
| Root MSE | 3778.94878 | R-square | 0.6951 | | |
| Dep Mean | 15463.00000 | Adj R-sq | 0.6935 | | |
| C.V. | 24.43865 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|-----------------------|-----------|
| INTERCEP | 1 | -21570 | 1823.8601874 | -11.826 | 0.0001 |
| TIME | 1 | 0.253081 | 0.09118303 | 2.776 | 0.0056 |
| MP | 1 | -1824.539417 | 309.65156086 | -5.892 | 0.0001 |
| FAMILYINC | 1 | 0.295782 | 0.01231313 | 24.022 | 0.0001 |
| TEMP | 1 | 193.357955 | 8.32224752 | 23.234 | 0.0001 |
| DAYS | 1 | 274.499875 | 49.93061661 | 5.498 | 0.0001 |
| COMPROXY | 1 | 6504.072456 | 1632.0700390 | 3.985 | 0.0001 |
| D11 | 1 | 1890.103558 | 285.36042013 | 6.624 | 0.0001 |
| DSEAS | 1 | 5919.128716 | 355.52834295 | 16.649 | 0.0001 |

Model: METROPLEX SURBURBAN: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 8 | 172.54513 | 21.56814 | 490.170 | 0.0000 |
| Error | 1443 | 63.49395 | 0.04400 | | |
| C Total | 1451 | 236.03907 | | | |
| Root MSE | 0.20976 | R-square | 0.7310 | | |
| Dep Mean | 9.56174 | Adj R-sq | 0.7295 | | |
| C.V. | 2.19380 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 7.144859 | 0.10124033 | 70.573 | 0.0000 |
| TIME | 1 | 0.000021755 | 0.00000506 | 4.298 | 0.0001 |
| MP | 1 | -0.137026 | 0.01718839 | -7.972 | 0.0001 |
| FAMLYINC | 1 | 0.000019751 | 0.00000068 | 28.897 | 0.0001 |
| TEMP | 1 | 0.013067 | 0.00046196 | 28.286 | 0.0001 |
| DAYS | 1 | 0.016637 | 0.00277159 | 6.003 | 0.0001 |
| COMPROXY | 1 | 0.341381 | 0.09059428 | 3.768 | 0.0002 |
| D11 | 1 | 0.112936 | 0.01584002 | 7.130 | 0.0001 |
| DSEAS | 1 | 0.251443 | 0.01973496 | 12.741 | 0.0001 |

Model: METROPLEX SUBURBAN: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 8 | 172.79151 | 21.59894 | 492.782 | 0.0000 |
| Error | 1443 | 63.24756 | 0.04383 | | |
| C Total | 1451 | 236.03907 | | | |
| Root MSE | 0.20936 | R-square | 0.7320 | | |
| Dep Mean | 9.56174 | Adj R-sq | 0.7306 | | |
| C.V. | 2.18954 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -5.139531 | 0.58384730 | -8.803 | 0.0001 |
| LOGTIME | 1 | 0.135581 | 0.04235433 | 3.201 | 0.0014 |
| LOGMP | 1 | -0.124848 | 0.02225116 | -5.611 | 0.0001 |
| LOGFINC | 1 | 0.842886 | 0.03061132 | 27.535 | 0.0001 |
| LOGTEMP | 1 | 0.734592 | 0.02697762 | 27.230 | 0.0001 |
| LOGDAYS | 1 | 0.467252 | 0.07352911 | 6.355 | 0.0001 |
| LOGPROXY | 1 | 0.101949 | 0.00970706 | 10.503 | 0.0001 |
| D11 | 1 | 0.078444 | 0.01579085 | 4.968 | 0.0001 |
| DSEAS | 1 | 0.305530 | 0.01876254 | 16.284 | 0.0001 |

METROPLEX SUBURBAN

CORRELATION ANALYSIS

6 'VAR' Variables: TIME MP FAMLYINC TEMP DAYS COMPROXY

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|------|----------|----------|-----------|
| TIME | 1452 | 8568 | 1160 | 12440472 |
| MP | 1452 | 1.32790 | 0.38287 | 1928 |
| FAMLYINC | 1452 | 48482 | 10777 | 70396490 |
| TEMP | 1452 | 65.17119 | 15.16228 | 94629 |
| DAYS | 1452 | 27.54821 | 2.15550 | 40000 |
| COMPROXY | 1452 | 0.09139 | 0.06618 | 132.69237 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| TIME | 6575 | 10562 |
| MP | 0.64286 | 2.30438 |
| FAMLYINC | 26312 | 87883 |
| TEMP | 32.58500 | 91.98500 |
| DAYS | 19.00000 | 31.00000 |
| COMPROXY | 0.01000 | 0.47917 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0 / N = 1452

| | TIME | MP | FAMLYINC | TEMP | DAYS | COMPROXY |
|----------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| TIME | 1.00000 0.0 | 0.30441 0.0001 | 0.00895 0.7332 | 0.04390 0.0945 | -0.03705 0.1582 | 0.00738 0.7788 |
| MP | 0.30441 0.0001 | 1.00000 0.0 | -0.16430 0.0001 | -0.09208 0.0004 | -0.04319 0.0999 | -0.08983 0.0006 |
| FAMLYINC | 0.00895 0.7332 | -0.16430 0.0001 | 1.00000 0.0 | -0.01140 0.6642 | -0.06932 0.0082 | 0.37098 0.0001 |
| TEMP | 0.04390 0.0945 | -0.09208 0.0004 | -0.01140 0.6642 | 1.00000 0.0 | 0.09785 0.0002 | -0.01485 0.5717 |
| DAYS | -0.03705 0.1582 | -0.04319 0.0999 | -0.06932 0.0082 | 0.09785 0.0002 | 1.00000 0.0 | -0.02979 0.2567 |
| COMPROXY | 0.00738 0.7788 | -0.08983 0.0006 | 0.37098 0.0001 | -0.01485 0.5717 | -0.02979 0.2567 | 1.00000 0.0 |

Model: ROLLING PLAINS REGION: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-------------|----------------|--------------|---------|--------|
| Model | 10 | 26430705741 | 2643070574.1 | 137.263 | 0.0001 |
| Error | 517 | 9955096869.0 | 19255506.516 | | |
| C Total | 527 | 36385802610 | | | |
| Root MSE | 4388.10967 | R-square | 0.7264 | | |
| Dep Mean | 21172.86364 | Adj R-sq | 0.7211 | | |
| C.V. | 20.72516 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -11913 | 6064.0404578 | -1.964 | 0.0500 |
| TIME | 1 | -0.776597 | 0.20586188 | -3.772 | 0.0002 |
| MP | 1 | -5931.982953 | 1269.1684474 | -4.674 | 0.0001 |
| FAMLYINC | 1 | 0.547395 | 0.19111070 | 2.864 | 0.0043 |
| TEMP | 1 | 286.207899 | 16.33152561 | 17.525 | 0.0001 |
| DAYS | 1 | 200.609337 | 97.19864266 | 2.064 | 0.0395 |
| COMPROXY | 1 | 33642 | 16283.981798 | 2.066 | 0.0393 |
| D1 | 1 | -9059.204931 | 2573.5042314 | -3.520 | 0.0005 |
| D21 | 1 | -7901.579490 | 2367.1474687 | -3.338 | 0.0009 |
| D28 | 1 | -9643 | 2826.3196940 | -3.412 | 0.0007 |
| DSEAS | 1 | 5345.606923 | 635.77889602 | 8.408 | 0.0001 |

Model: ROLLING PLAINS REGION: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 10 | 61.40194 | 6.14019 | 136.704 | 0.0001 |
| Error | 517 | 23.22164 | 0.04492 | | |
| C Total | 527 | 84.62358 | | | |
| Root MSE | 0.21193 | R-square | 0.7256 | | |
| Dep Mean | 9.88248 | Adj R-sq | 0.7203 | | |
| C.V. | 2.14454 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|-----------------------|-----------|
| INTERCEP | 1 | 8.015805 | 0.29287736 | 27.369 | 0.0001 |
| TIME | 1 | -0.000028619 | 0.00000994 | -2.878 | 0.0042 |
| MP | 1 | -0.299427 | 0.06129753 | -4.885 | 0.0001 |
| FAMLYINC | 1 | 0.000030872 | 0.00000923 | 3.345 | 0.0009 |
| TEMP | 1 | 0.015248 | 0.00078877 | 19.332 | 0.0001 |
| DAYS | 1 | 0.009884 | 0.00469444 | 2.105 | 0.0357 |
| COMPROXY | 1 | 1.354971 | 0.78647391 | 1.723 | 0.0855 |
| D1 | 1 | -0.487448 | 0.12429355 | -3.922 | 0.0001 |
| D21 | 1 | -0.418419 | 0.11432706 | -3.660 | 0.0003 |
| D28 | 1 | -0.521871 | 0.13650388 | -3.823 | 0.0001 |
| DSEAS | 1 | 0.172291 | 0.03070646 | 5.611 | 0.0001 |

Model: ROLLING PLAINS REGION: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 10 | 60.72633 | 6.07263 | 131.377 | 0.0001 |
| Error | 517 | 23.89725 | 0.04622 | | |
| C Total | 527 | 84.62358 | | | |
| Root MSE | 0.21500 | R-square | 0.7176 | | |
| Dep Mean | 9.88248 | Adj R-sq | 0.7121 | | |
| C.V. | 2.17552 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -6.442416 | 3.87235629 | -1.664 | 0.0968 |
| TIME | 1 | -0.000029889 | 0.00000988 | -3.025 | 0.0026 |
| LOGMP | 1 | -0.275934 | 0.05368670 | -5.140 | 0.0001 |
| LOGFINC | 1 | 1.198870 | 0.36294393 | 3.303 | 0.0010 |
| LOGTEMP | 1 | 0.871546 | 0.04696277 | 18.558 | 0.0001 |
| LOGDAYS | 1 | 0.233351 | 0.12088692 | 1.930 | 0.0541 |
| LOGPROXY | 1 | 0.122509 | 0.06121544 | 2.001 | 0.0459 |
| D1 | 1 | -0.500443 | 0.12209892 | -4.099 | 0.0001 |
| D21 | 1 | -0.427581 | 0.11529756 | -3.708 | 0.0002 |
| D28 | 1 | -0.518969 | 0.13425101 | -3.866 | 0.0001 |
| DSEAS | 1 | 0.224957 | 0.02981238 | 7.546 | 0.0001 |

ROLLING PLAINS REGION

CORRELATION ANALYSIS

6 'VAR' Variables: TIME MP FAMLYINC TEMP DAYS COMPROXY

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|-----|----------|----------|-----------|
| TIME | 576 | 8385 | 1266 | 4829880 |
| MP | 528 | 0.79049 | 0.19714 | 417.37956 |
| FAMLYINC | 576 | 40442 | 4822 | 23294502 |
| TEMP | 576 | 66.27068 | 15.08794 | 38172 |
| DAYS | 576 | 27.96528 | 2.01529 | 16108 |
| COMPROXY | 576 | 0.09876 | 0.03096 | 56.88847 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| TIME | 6210 | 10562 |
| MP | 0.52321 | 1.36442 |
| FAMLYINC | 30353 | 45445 |
| TEMP | 30.08000 | 91.90500 |
| DAYS | 13.00000 | 31.00000 |
| COMPROXY | 0.03000 | 0.16000 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0
/ Number of Observations

| | TIME | MP | FAMLYINC | TEMP | DAYS | COMPROXY |
|----------|---------------------------|---------------------------|---------------------------|---------------------------|---------------------------|---------------------------|
| TIME | 1.00000 0.0 576 | 0.10292 0.0180 528 | 0.19152 0.0001 576 | 0.01078 0.7962 576 | -0.04857 0.2445 576 | 0.38428 0.0001 576 |
| MP | 0.10292 0.0180 528 | 1.00000 0.0 528 | 0.58620 0.0001 528 | -0.18135 0.0001 528 | -0.11763 0.0068 528 | -0.38412 0.0001 528 |
| FAMLYINC | 0.19152 0.0001 576 | 0.58620 0.0001 528 | 1.00000 0.0 576 | -0.25169 0.0001 576 | -0.13207 0.0015 576 | -0.32234 0.0001 576 |
| TEMP | 0.01078 0.7962 576 | -0.18135 0.0001 528 | -0.25169 0.0001 576 | 1.00000 0.0 576 | -0.03190 0.4448 576 | 0.14894 0.0003 576 |
| DAYS | -0.04857 0.2445 576 | -0.11763 0.0068 528 | -0.13207 0.0015 576 | -0.03190 0.4448 576 | 1.00000 0.0 576 | -0.02403 0.5649 576 |
| COMPROXY | 0.38428 0.0001 576 | -0.38412 0.0001 528 | -0.32234 0.0001 576 | 0.14894 0.0003 576 | -0.02403 0.5649 576 | 1.00000 0.0 576 |

Model: SOUTHEAST REGION: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-------------|----------------|--------------|---------|--------|
| Model | 16 | 26298551955 | 1643659497.2 | 311.859 | 0.0000 |
| Error | 1423 | 7499954375.5 | 5270523.1029 | | |
| C Total | 1439 | 33798506331 | | | |
| Root MSE | 2295.76199 | R-square | 0.7781 | | |
| Dep Mean | 15434.25417 | Adj R-sq | 0.7756 | | |
| C.V. | 14.87446 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|-----------------------|-----------|
| INTERCEP | 1 | -1996.386490 | 1154.4456417 | -1.729 | 0.0840 |
| TIME | 1 | -0.181954 | 0.06163899 | -2.952 | 0.0032 |
| MP | 1 | -1395.244115 | 236.28005589 | -5.905 | 0.0001 |
| FAMILYINC | 1 | 0.237526 | 0.01331346 | 17.841 | 0.0001 |
| TEMP | 1 | 86.421350 | 6.30083100 | 13.716 | 0.0001 |
| DAYS | 1 | 96.150638 | 25.70203944 | 3.741 | 0.0002 |
| COMPROXY | 1 | 4990.709171 | 1262.0879326 | 3.954 | 0.0001 |
| D8 | 1 | 6408.858508 | 262.22546199 | 24.440 | 0.0001 |
| GAL | 1 | 4732.750454 | 276.68752336 | 17.105 | 0.0001 |
| LEA | 1 | -1648.659662 | 339.59336569 | -4.855 | 0.0001 |
| TEX | 1 | -1009.821740 | 259.24902778 | -3.895 | 0.0001 |
| ALV | 1 | -1409.286176 | 253.19825131 | -5.566 | 0.0001 |
| ANG | 1 | -2687.339961 | 252.73317654 | -10.633 | 0.0001 |
| BRA | 1 | -4057.631951 | 297.05337000 | -13.660 | 0.0001 |
| FRE | 1 | 5035.064919 | 313.38021464 | 16.067 | 0.0001 |
| LAK | 1 | -2423.920637 | 252.37558931 | -9.604 | 0.0001 |
| DSEAS | 1 | 1349.493424 | 195.41129353 | 6.906 | 0.0001 |

Model: SOUTHEAST REGION: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 16 | 115.17673 | 7.19855 | 376.446 | 0.0000 |
| Error | 1423 | 27.21116 | 0.01912 | | |
| C Total | 1439 | 142.38789 | | | |
| Root MSE | 0.13828 | R-square | 0.8089 | | |
| Dep Mean | 9.59558 | Adj R-sq | 0.8067 | | |
| C.V. | 1.44112 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 8.493204 | 0.06953729 | 122.139 | 0.0000 |
| TIME | 1 | -0.000013442 | 0.00000371 | -3.620 | 0.0003 |
| MP | 1 | -0.069661 | 0.01423218 | -4.895 | 0.0001 |
| FAMILYINC | 1 | 0.000015305 | 0.00000080 | 19.085 | 0.0001 |
| TEMP | 1 | 0.005717 | 0.00037953 | 15.064 | 0.0001 |
| DAYS | 1 | 0.006189 | 0.00154815 | 3.998 | 0.0001 |
| COMPROXY | 1 | 0.237329 | 0.07602105 | 3.122 | 0.0018 |
| D8 | 1 | 0.340880 | 0.01579498 | 21.582 | 0.0001 |
| GAL | 1 | 0.276078 | 0.01666609 | 16.565 | 0.0001 |
| LEA | 1 | -0.138328 | 0.02045519 | -6.762 | 0.0001 |
| TEX | 1 | -0.060800 | 0.01561570 | -3.894 | 0.0001 |
| ALV | 1 | -0.094087 | 0.01525123 | -6.169 | 0.0001 |
| ANG | 1 | -0.190705 | 0.01522322 | -12.527 | 0.0001 |
| BRA | 1 | -0.397627 | 0.01789282 | -22.223 | 0.0001 |
| FRE | 1 | 0.286827 | 0.01887625 | 15.195 | 0.0001 |
| LAK | 1 | -0.178845 | 0.01520168 | -11.765 | 0.0001 |
| DSEAS | 1 | 0.068892 | 0.01177047 | 5.853 | 0.0001 |

Model: SOUTHEAST REGION: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|------|----------------|-------------|---------|--------|
| Model | 16 | 114.96438 | 7.18527 | 372.842 | 0.0000 |
| Error | 1423 | 27.42351 | 0.01927 | | |
| C Total | 1439 | 142.38789 | | | |
| Root MSE | | 0.13882 | R-square | 0.8074 | |
| Dep Mean | | 9.59558 | Adj R-sq | 0.8052 | |
| C.V. | | 1.44673 | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 1.389987 | 0.61790246 | 2.250 | 0.0246 |
| LOGTIME | 1 | -0.152247 | 0.03095027 | -4.919 | 0.0001 |
| LOGMP | 1 | -0.087162 | 0.02005363 | -4.346 | 0.0001 |
| LOGFINC | 1 | 0.718286 | 0.04094729 | 17.542 | 0.0001 |
| LOGTEMP | 1 | 0.354680 | 0.02438446 | 14.545 | 0.0001 |
| LOGDAYS | 1 | 0.142706 | 0.04028980 | 3.542 | 0.0004 |
| LOGPROXY | 1 | 0.037680 | 0.00966490 | 3.899 | 0.0001 |
| D8 | 1 | 0.330731 | 0.01584576 | 20.872 | 0.0001 |
| GAL | 1 | 0.280520 | 0.01695594 | 16.544 | 0.0001 |
| LEA | 1 | -0.145589 | 0.01979059 | -7.356 | 0.0001 |
| TEX | 1 | -0.044898 | 0.01740176 | -2.580 | 0.0100 |
| ALV | 1 | -0.086214 | 0.01532686 | -5.625 | 0.0001 |
| ANG | 1 | -0.195783 | 0.01536812 | -12.740 | 0.0001 |
| BRA | 1 | -0.380608 | 0.01830996 | -20.787 | 0.0001 |
| FRE | 1 | 0.297972 | 0.01714965 | 17.375 | 0.0001 |
| LAK | 1 | -0.179347 | 0.01533369 | -11.696 | 0.0001 |
| DSEAS | 1 | 0.080547 | 0.01149926 | 7.005 | 0.0001 |

SOUTHEAST REGION

CORRELATION ANALYSIS

6 'VAR' Variables: TIME MP FAMLYINC TEMP DAYS COMPROXY

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|------|----------|----------|-----------|
| TIME | 1440 | 8565 | 1165 | 12333272 |
| MP | 1440 | 1.32836 | 0.45820 | 1913 |
| FAMLYINC | 1440 | 47817 | 6993 | 68857159 |
| TEMP | 1440 | 68.78029 | 11.64441 | 99044 |
| DAYS | 1440 | 26.77361 | 2.41955 | 38554 |
| COMPROXY | 1440 | 0.10522 | 0.07447 | 151.52273 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| TIME | 6575 | 10562 |
| MP | 0.65268 | 3.28659 |
| FAMLYINC | 30266 | 65054 |
| TEMP | 43.13000 | 87.87000 |
| DAYS | 18.00000 | 31.00000 |
| COMPROXY | 0.02020 | 0.40404 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0 / N = 1440

| | TIME | MP | FAMLYINC | TEMP | DAYS | COMPROXY |
|----------|--------------------|-------------------|--------------------|--------------------|--------------------|--------------------|
| TIME | 1.00000 0.0 | 0.29621 0.0001 | -0.07862 0.0028 | 0.06388 0.0153 | 0.05639 0.0324 | 0.17312 0.0001 |
| MP | 0.29621 0.0001 | 1.00000 0.0 | 0.16675 0.0001 | 0.00753 0.7754 | 0.01296 0.6232 | 0.28966 0.0001 |
| FAMLYINC | -0.07862 0.0028 | 0.16675 0.0001 | 1.00000 0.0 | -0.01030 0.6960 | -0.01305 0.6206 | 0.48177 0.0001 |
| TEMP | 0.06388 0.0153 | 0.00753 0.7754 | -0.01030 0.6960 | 1.00000 0.0 | -0.00211 0.9362 | 0.01014 0.7005 |
| DAYS | 0.05639 0.0324 | 0.01296 0.6232 | -0.01305 0.6206 | -0.00211 0.9362 | 1.00000 0.0 | -0.01926 0.4653 |
| COMPROXY | 0.17312 0.0001 | 0.28966 0.0001 | 0.48177 0.0001 | 0.01014 0.7005 | -0.01926 0.4653 | 1.00000 0.0 |

Model: VALLEY REGION: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-------------|----------------|--------------|---------|--------|
| Model | 11 | 23350872225 | 2122806565.9 | 85.118 | 0.0001 |
| Error | 912 | 22745023776 | 24939719.053 | | |
| C Total | 923 | 46095896002 | | | |
| Root MSE | 4993.96827 | R-square | 0.5066 | | |
| Dep Mean | 22954.61147 | Adj R-sq | 0.5006 | | |
| C.V. | 21.75584 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|-----------------------|-----------|
| INTERCEP | 1 | 5505.802596 | 3664.6024603 | 1.502 | 0.1333 |
| TIME | 1 | -0.829722 | 0.16237484 | -5.110 | 0.0001 |
| MP | 1 | -8656.641869 | 1264.4735182 | -6.846 | 0.0001 |
| FAMILYINC | 1 | 0.220667 | 0.04318008 | 5.110 | 0.0001 |
| TEMP | 1 | 203.516656 | 19.09267294 | 10.659 | 0.0001 |
| DAYS | 1 | 409.280531 | 86.11968404 | 4.752 | 0.0001 |
| COMPROXY | 1 | 24930 | 4142.5910635 | 6.018 | 0.0001 |
| MCA | 1 | 3543.464684 | 522.40251895 | 6.783 | 0.0001 |
| EDI | 1 | 1856.729183 | 623.37364733 | 2.979 | 0.0030 |
| PHA | 1 | 905.216465 | 691.43213718 | 1.309 | 0.1908 |
| HAR | 1 | -6967.380343 | 665.20517369 | -10.474 | 0.0001 |
| DSEAS | 1 | 3145.192855 | 522.37571337 | 6.021 | 0.0001 |

Model: VALLEY REGION: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 11 | 45.55071 | 4.14097 | 92.801 | 0.0001 |
| Error | 912 | 40.69554 | 0.04462 | | |
| C Total | 923 | 86.24625 | | | |
| Root MSE | 0.21124 | R-square | 0.5281 | | |
| Dep Mean | 9.99500 | Adj R-sq | 0.5225 | | |
| C.V. | 2.11346 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 9.211863 | 0.15500906 | 59.428 | 0.0000 |
| TIME | 1 | -0.000036579 | 0.00000687 | -5.326 | 0.0001 |
| MP | 1 | -0.411456 | 0.05348598 | -7.693 | 0.0001 |
| FAMILYINC | 1 | 0.000010704 | 0.00000183 | 5.861 | 0.0001 |
| TEMP | 1 | 0.009106 | 0.00080760 | 11.276 | 0.0001 |
| DAYS | 1 | 0.017671 | 0.00364278 | 4.851 | 0.0001 |
| COMPROXY | 1 | 1.042447 | 0.17522751 | 5.949 | 0.0001 |
| MCA | 1 | 0.180689 | 0.02209711 | 8.177 | 0.0001 |
| EDI | 1 | 0.112436 | 0.02636809 | 4.264 | 0.0001 |
| PHA | 1 | 0.072082 | 0.02924690 | 2.465 | 0.0139 |
| HAR | 1 | -0.299632 | 0.02813752 | -10.649 | 0.0001 |
| DSEAS | 1 | 0.115354 | 0.02209598 | 5.221 | 0.0001 |

Model: VALLEY REGION: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-----|----------------|-------------|---------|--------|
| Model | 11 | 45.03913 | 4.09447 | 90.619 | 0.0001 |
| Error | 912 | 41.20712 | 0.04518 | | |
| C Total | 923 | 86.24625 | | | |
| Root MSE | | 0.21256 | R-square | 0.5222 | |
| Dep Mean | | 9.99500 | Adj R-sq | 0.5165 | |
| C.V. | | 2.12670 | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 8.961255 | 0.73180172 | 12.245 | 0.0001 |
| LOGTIME | 1 | -0.340250 | 0.05561928 | -6.117 | 0.0001 |
| LOGMP | 1 | -0.389120 | 0.04188872 | -9.289 | 0.0001 |
| LOGFINC | 1 | 0.055281 | 0.02653176 | 2.084 | 0.0375 |
| LOGTEMP | 1 | 0.617653 | 0.05604660 | 11.020 | 0.0001 |
| LOGDAYS | 1 | 0.414961 | 0.09606856 | 4.319 | 0.0001 |
| LOGPROXY | 1 | 0.141904 | 0.01713603 | 8.281 | 0.0001 |
| MCA | 1 | 0.159901 | 0.02238454 | 7.143 | 0.0001 |
| EDI | 1 | 0.162060 | 0.02483033 | 6.527 | 0.0001 |
| PHA | 1 | 0.142604 | 0.02746259 | 5.193 | 0.0001 |
| HAR | 1 | -0.397773 | 0.02691119 | -14.781 | 0.0001 |
| DSEAS | 1 | 0.126584 | 0.02175719 | 5.818 | 0.0001 |

VALLEY REGION

CORRELATION ANALYSIS

6 'VAR' Variables: TIME MP FAMLYINC TEMP DAYS COMPROXY

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|-----|----------|----------|-----------|
| TIME | 924 | 8568 | 1160 | 7916664 |
| MP | 924 | 0.81986 | 0.20528 | 757.55452 |
| FAMLYINC | 924 | 31167 | 5216 | 28798617 |
| TEMP | 924 | 73.57111 | 10.31188 | 67980 |
| DAYS | 924 | 28.04870 | 1.93581 | 25917 |
| COMPROXY | 924 | 0.11519 | 0.05849 | 106.43547 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| TIME | 6575 | 10562 |
| MP | 0.46012 | 1.29199 |
| FAMLYINC | 2993 | 48955 |
| TEMP | 50.12500 | 89.61000 |
| DAYS | 16.00000 | 31.00000 |
| COMPROXY | 0.01010 | 0.36842 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0 / N = 924

| | TIME | MP | FAMLYINC | TEMP | DAYS | COMPROXY |
|----------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| TIME | 1.00000 0.0 | 0.28738 0.0001 | -0.19077 0.0001 | 0.05093 0.1219 | 0.03006 0.3614 | 0.03664 0.2659 |
| MP | 0.28738 0.0001 | 1.00000 0.0 | 0.03932 0.2324 | 0.00483 0.8833 | -0.05041 0.1257 | -0.01003 0.7608 |
| FAMLYINC | -0.19077 0.0001 | 0.03932 0.2324 | 1.00000 0.0 | 0.00101 0.9756 | -0.00437 0.8944 | 0.30401 0.0001 |
| TEMP | 0.05093 0.1219 | 0.00483 0.8833 | 0.00101 0.9756 | 1.00000 0.0 | -0.12053 0.0002 | 0.00581 0.8600 |
| DAYS | 0.03006 0.3614 | -0.05041 0.1257 | -0.00437 0.8944 | -0.12053 0.0002 | 1.00000 0.0 | -0.00582 0.8598 |
| COMPROXY | 0.03664 0.2659 | -0.01003 0.7608 | 0.30401 0.0001 | 0.00581 0.8600 | -0.00582 0.8598 | 1.00000 0.0 |

Model: WEST REGION: LINEAR
 Dependent Variable: CONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|-----|----------------|--------------|---------|--------|
| Model | 6 | 31332251800 | 5222041966.7 | 297.255 | 0.0001 |
| Error | 653 | 11471626569 | 17567575.145 | | |
| C Total | 659 | 42803878370 | | | |
| Root MSE | | 4191.36913 | R-square | 0.7320 | |
| Dep Mean | | 21941.36818 | Adj R-sq | 0.7295 | |
| C.V. | | 19.10259 | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -36654 | 3412.3052204 | -10.742 | 0.0001 |
| MP | 1 | -1929.111210 | 717.35804521 | -2.689 | 0.0073 |
| FAMLYINC | 1 | 0.122150 | 0.02604647 | 4.690 | 0.0001 |
| TEMP | 1 | 382.317129 | 20.79378033 | 18.386 | 0.0001 |
| DAYS | 1 | 1021.504243 | 104.76109354 | 9.751 | 0.0001 |
| COMPROXY | 1 | 8231.467259 | 2180.7883385 | 3.775 | 0.0002 |
| DSEAS | 1 | 3141.149196 | 647.23688346 | 4.853 | 0.0001 |

Model: WEST REGION: LOG-LINEAR
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 6 | 63.99562 | 10.66594 | 335.751 | 0.0001 |
| Error | 653 | 20.74412 | 0.03177 | | |
| C Total | 659 | 84.73973 | | | |
| Root MSE | 0.17823 | R-square | 0.7552 | | |
| Dep Mean | 9.93168 | Adj R-sq | 0.7530 | | |
| C.V. | 1.79460 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|-----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | 7.381892 | 0.14510516 | 50.873 | 0.0001 |
| MP | 1 | -0.063118 | 0.03050500 | -2.069 | 0.0389 |
| FAMILYINC | 1 | 0.000005209 | 0.00000111 | 4.703 | 0.0001 |
| TEMP | 1 | 0.018515 | 0.00088424 | 20.938 | 0.0001 |
| DAYS | 1 | 0.040344 | 0.00445487 | 9.056 | 0.0001 |
| COMPROXY | 1 | 0.384444 | 0.09273603 | 4.146 | 0.0001 |
| DSEAS | 1 | 0.097948 | 0.02752316 | 3.559 | 0.0004 |

Model: WEST REGION: LOG-LOG
 Dependent Variable: LOGCONS

Analysis of Variance

| Source | DF | Sum of Squares | Mean Square | F Value | Prob>F |
|----------|---------|----------------|-------------|---------|--------|
| Model | 6 | 61.81668 | 10.30278 | 293.491 | 0.0001 |
| Error | 653 | 22.92305 | 0.03510 | | |
| C Total | 659 | 84.73973 | | | |
| Root MSE | 0.18736 | R-square | 0.7295 | | |
| Dep Mean | 9.93168 | Adj R-sq | 0.7270 | | |
| C.V. | 1.88650 | | | | |

Parameter Estimates

| Variable | DF | Parameter Estimate | Standard Error | T for H0: Parameter=0 | Prob > T |
|----------|----|--------------------|----------------|--------------------------|-----------|
| INTERCEP | 1 | -0.857688 | 0.76959397 | -1.114 | 0.2655 |
| LOGMP | 1 | -0.099418 | 0.03133803 | -3.172 | 0.0016 |
| LOGFINC | 1 | 0.305497 | 0.05608852 | 5.447 | 0.0001 |
| LOGTEMP | 1 | 0.866224 | 0.04760267 | 18.197 | 0.0001 |
| LOGDAYS | 1 | 1.187610 | 0.12995848 | 9.138 | 0.0001 |
| LOGPROXY | 1 | 0.040590 | 0.01467332 | 2.766 | 0.0058 |
| DSEAS | 1 | 0.205719 | 0.02587832 | 7.949 | 0.0001 |

WEST REGION

CORRELATION ANALYSIS

6 'VAR' Variables: MP FAMLYINC TEMP DAYS COMPROXY DSEAS

Simple Statistics

| Variable | N | Mean | Std Dev | Sum |
|----------|-----|----------|----------|-----------|
| MP | 660 | 0.97979 | 0.29601 | 646.66131 |
| FAMLYINC | 720 | 46401 | 8231 | 33408653 |
| TEMP | 720 | 61.31661 | 15.07809 | 44148 |
| DAYS | 720 | 28.60139 | 1.67057 | 20593 |
| COMPROXY | 720 | 0.11817 | 0.07649 | 85.08445 |
| DSEAS | 720 | 0.41667 | 0.49335 | 300.00000 |

Simple Statistics

| Variable | Minimum | Maximum |
|----------|----------|----------|
| MP | 0.38008 | 1.83861 |
| FAMLYINC | 35396 | 61683 |
| TEMP | 24.71000 | 87.14500 |
| DAYS | 22.00000 | 31.00000 |
| COMPROXY | 0.04167 | 0.41000 |
| DSEAS | 0 | 1.00000 |

Pearson Correlation Coefficients / Prob > |R| under Ho: Rho=0 / Number of Observations

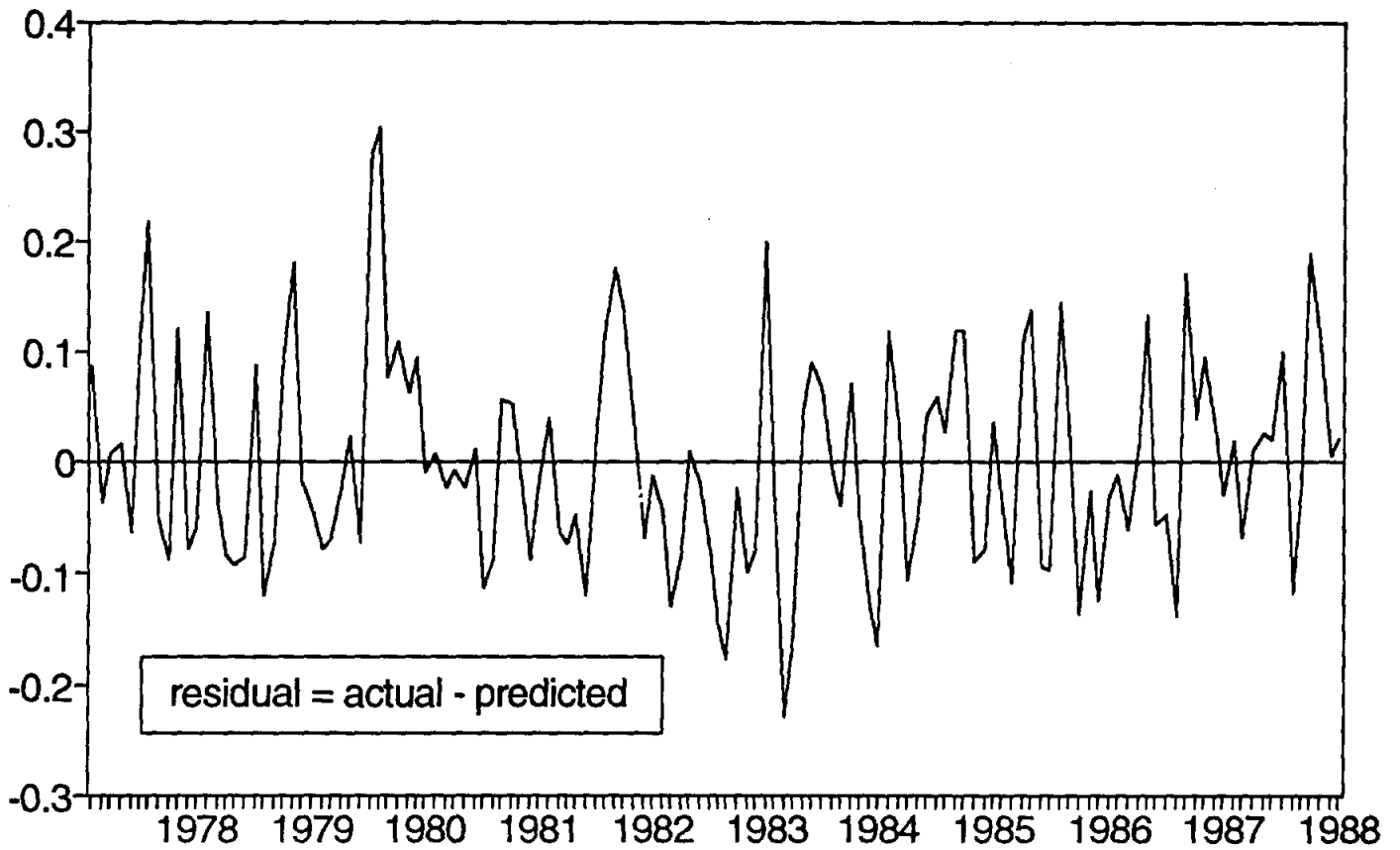
| | MP | FAMLYINC | TEMP | DAYS | COMPROXY | DSEAS |
|----------|---------------------------|---------------------------|---------------------------|---------------------------|---------------------------|---------------------------|
| MP | 1.00000 0.0 660 | 0.63148 0.0001 660 | -0.00873 0.8229 660 | -0.11917 0.0022 660 | 0.19112 0.0001 660 | -0.04058 0.2979 660 |
| FAMLYINC | 0.63148 0.0001 660 | 1.00000 0.0 720 | 0.07132 0.0558 720 | -0.12847 0.0005 720 | 0.16096 0.0001 720 | 0.00000 1.0000 720 |
| TEMP | -0.00873 0.8229 660 | 0.07132 0.0558 720 | 1.00000 0.0 720 | -0.25119 0.0001 720 | -0.06139 0.0998 720 | 0.84892 0.0001 720 |
| DAYS | -0.11917 0.0022 660 | -0.12847 0.0005 720 | -0.25119 0.0001 720 | 1.00000 0.0 720 | -0.11694 0.0017 720 | -0.32640 0.0001 720 |
| COMPROXY | 0.19112 0.0001 660 | 0.16096 0.0001 720 | -0.06139 0.0998 720 | -0.11694 0.0017 720 | 1.00000 0.0 720 | 0.00000 1.0000 720 |
| DSEAS | -0.04058 0.2979 660 | 0.00000 1.0000 720 | 0.84892 0.0001 720 | -0.32640 0.0001 720 | 0.00000 1.0000 720 | 1.00000 0.0 720 |

APPENDIX B

Residuals Comparisons for Selected Cities

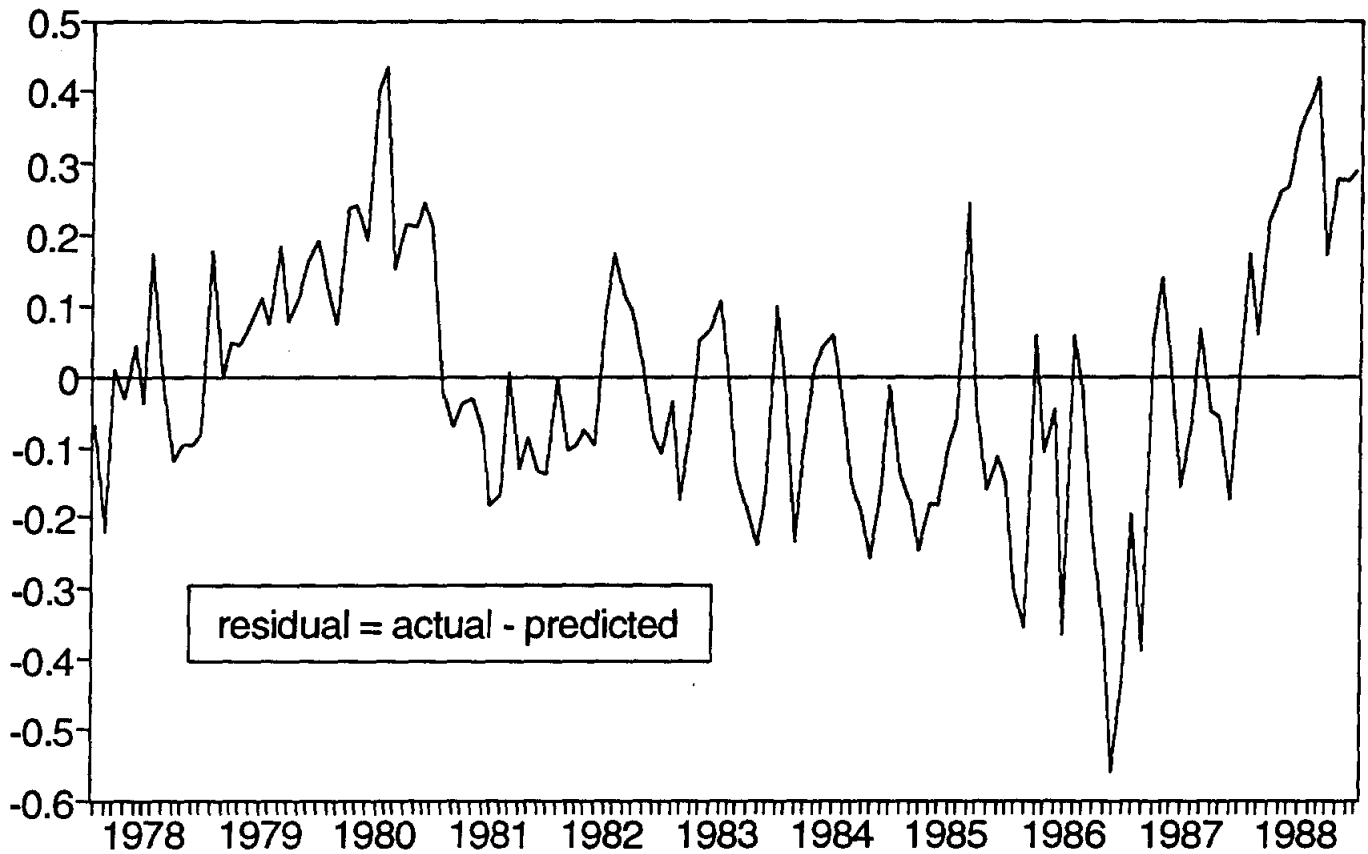
CITY OF AUSTIN

residuals of log of consumption



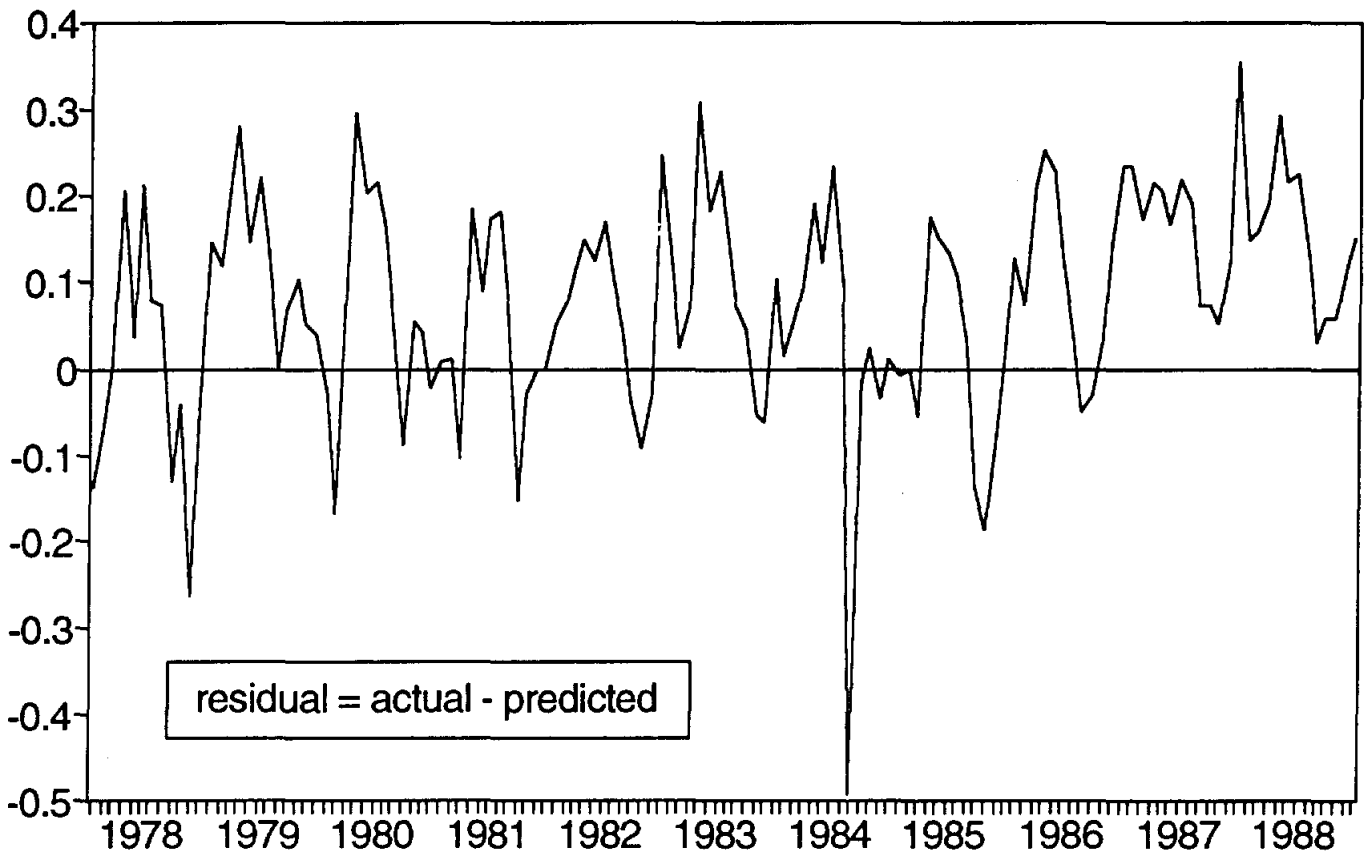
CITY OF CORPUS CHRISTI

residuals of log of consumption



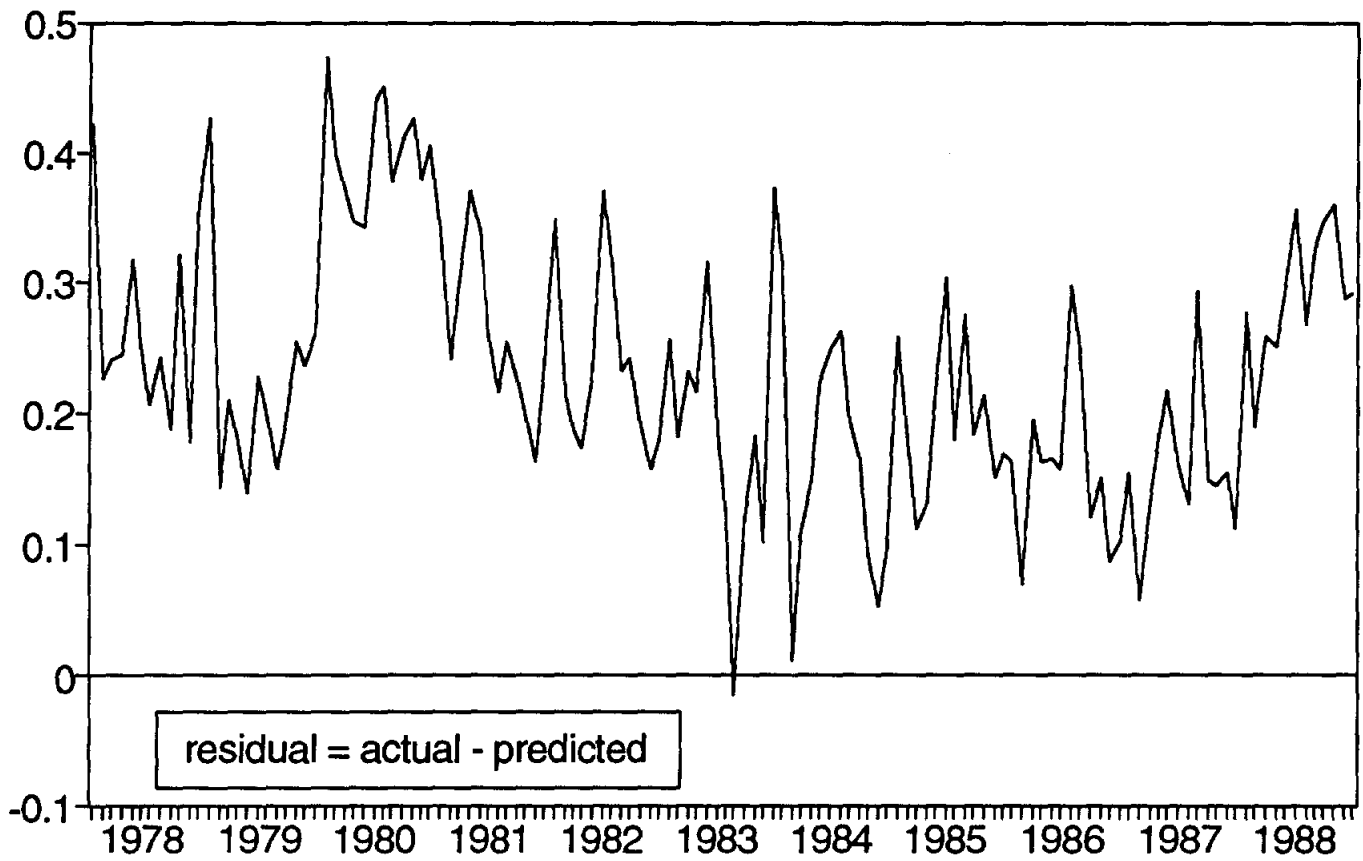
CITY OF EL PASO

residuals of log of consumption



CITY OF GALVESTON

residuals of log of consumption



CITY OF LAKE JACKSON

residuals of log of consumption

